

PONTIFÍCIA UNIVERSIDADE CATÓLICA DO PARANÁ  
CENTRO DE CIÊNCIAS DA SAÚDE  
PÓS GRADUAÇÃO EM CIÊNCIAS DA SAÚDE

**CRISTINA PELLEGRINO BAENA**

**DOENÇA CARDIOVASCULAR :  
TENDÊNCIA DE MORTALIDADE NO BRASIL E PREVENÇÃO  
GLOBAL**

CURITIBA/2013

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Tese apresentada ao Programa de Pós Graduação em Ciências da Saúde PPGCS da Escola de Medicina na Universidade Católica do Paraná PUCPR como requisito parcial de obtenção do título de Doutora.

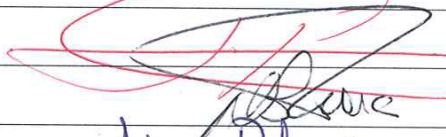
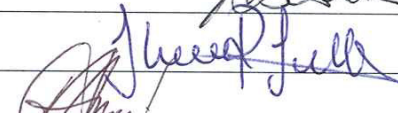
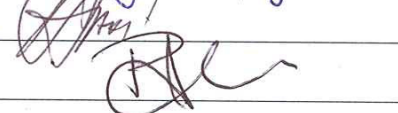
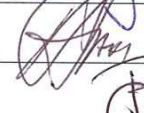

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**ATA DA SESSÃO PÚBLICA DE EXAME DE TESE DO PROGRAMA DE PÓS-GRADUAÇÃO EM CIÊNCIAS DA SAÚDE EM NÍVEL DE DOUTORADO DA PONTIFÍCIA UNIVERSIDADE CATÓLICA DO PARANÁ**

Ao dez dias do mês de abril de 2013, realizou-se a sessão pública de defesa de tese “DOENÇA CARDIOVASCULAR: TENDÊNCIA DE MORTALIDADE NO BRASIL E PREVENÇÃO GLOBAL” apresentada por **CRISTINA PELLEGRINO BAENA** para obtenção do título de doutor; Área de Concentração: Medicina e áreas afins.

A Banca Examinadora foi composta pelos seguintes membros:

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Prof. Dr. Paulo Andrade Lotufo (USP)	

De acordo com as normas regimentais a Banca Examinadora deliberou sobre os conceitos a serem distribuídos e que foram os seguintes:

Prof. Dr. José Rocha Faria Neto Parecer: Aprovado  
 Prof. Dr. Dalton Bertolim Précoma Parecer: Aprovado  
 Profª. Drª. Karin Regina Luhm Parecer: Aprovada  
 Prof. Dr. Andrei Carvalho Sposito Parecer: Aprovado  
 Prof. Dr. Paulo Andrade Lotufo Parecer: Aprovado  
**Parecer Final: APROVADA**


Observações da Banca Examinadora:

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Prof. Dr. José Rocha Faria Neto  
Presidente da Banca Examinadora

  
Prof. Dr. Roberto Pecoits Filho  
Coordenador do PPGCS PUCPR

***“If the major determinants of health are social, so must be the remedies.”***

***Prof.Sir M.Marmot***

## AGRADECIMENTOS

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## **LISTA DE ABREVIATURAS**

AMI – Acute Myocardial Infarction

ASMR – Age Standardized Mortality Rate

IAM- Infarto Agudo do Miocárdio

ICD – International Code of Diseases

CVD – Cardiovascular Diseases

DCV – Doença Cardiovascular

DCbV- Doença Cerebrovascular

DIC – Doença Isquêmica Cardíaca

ES – Effect Size

IHD – Ischemic Heart Diseases

CbVD – Cerebrovascular Diseases

OMS – Organização Mundial de Saúde

SD – Standard Deviation

SE – Standard Error

CI – Confidence Interval

WHO – World Health Organization

WMD – Weighted Mean Difference

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## RESUMO

**Introdução:** A doença cardiovascular (DCV) é a maior causa de óbitos nos países em desenvolvimento, e o Brasil não é exceção. Adicionalmente, a escalada de fatores de risco modificáveis como a pressão arterial elevada responde pelo impacto das doenças cardiovasculares em nível local e global. **Métodos:** Para as análises de mortalidade, dados demográficos foram obtidos do Instituto Brasileiro de Geografia e Estatística e dados de óbitos foram obtidos no Sistema de Informação de Mortalidade do Ministério da Saúde considerando gênero, faixa etária e local de residência. Em nível local, a causa de óbito estudada foi Infarto Agudo do Miocárdio (CID I-21) e foi construído um modelo de Regressão de Poisson. Foram estimadas taxas de mortalidade e número de mortes esperadas e não observadas entre 1999 a 2009. Em nível nacional, considerou-se como causas de óbito as Doenças Isquêmica Cardíacas (CID I20 a I25). Foi incluída toda a população brasileira ( $\geq 20$  anos), das cinco regiões entre 2000 a 2010. Foram calculadas as taxas padronizadas específicas por idade por 100.000 habitantes por sexo e região, tendo a população brasileira do mesmo ano e idade como padrão. Foram construídos modelos de regressão múltipla para quantificar as tendências temporais e projetar futuras taxas de óbito até 2015. Para a meta análise sobre intervenções não farmacológicas de controle e tratamento de pressão arterial elevada em populações de países em desenvolvimento, foi realizada uma revisão sistemática de ensaios clínicos publicados até setembro/2012 que relatassem efeitos na pressão arterial de qualquer intervenção relacionada a estilo de vida nestas populações. As intervenções foram agrupadas como educacionais, modificações de dieta, atividade física e combinadas. Foram calculados os efeitos das intervenções em mm de Hg (diferença média ponderada entre pressão arterial pré e pós intervenção) com o modelo de efeito *random*. Foram ainda realizadas análise de subgrupo e meta regressão para analisar as causas de heterogeneidade. **Resultados:** Em Curitiba, a tendência de óbito por IAM apresentou declínio significativo ( $p < 0,001$ ) no período com redução média anual de 3,8% (IC95%: 3,2% - 4,5%) sem diferença entre os gêneros ( $p = 0,238$ ), entretanto, as taxas padronizadas específicas por idade diferiram significativamente entre as faixas etárias ( $p = 0,018$ ). No Brasil, 560.710 homens e 406.627 mulheres morreram por DIC entre 2000 e 2010. A tendência de mortalidade por DIC apresentou-se estagnada, sendo influenciada pelas tendências convergentes de declínio no Sul e Sudeste e elevação no Norte e Nordeste ( $p < 0,05$ ) para ambos os gêneros. As projeções apontam para uma ampliação das discrepâncias entre Norte e Sul nos próximos anos. Por meio da revisão sistemática, foram identificados 6.211 publicações de intervenções não farmacológicas sobre a pressão arterial de populações de países em desenvolvimento, dos quais 43 estudos foram incluídos na meta análise compreendendo 6.779 participantes não repetidos. O efeito (IC 95%) calculado em mm Hg para pressão arterial sistólica foram -5.39 (-10.73, -0.05) em intervenções educacionais, -3.48 (-5.45, -1.50) para modificações de dieta, -11.37 (-16.06, -6.68) para atividade física e -6.09 (-8.87, -3.32) para intervenções combinadas. Os resultados foram semelhantes para pressão arterial diastólica. Entretanto, a heterogeneidade entre os estudos foi alta e explicada em grande parte ao tamanho das amostras, duração das intervenções e tipo de análise utilizada. **Conclusão:** A tendência de mortalidade por IAM apresentou declínio significativo em Curitiba na última década. No Brasil, a tendência de mortalidade por DIC que apresentava declínio, passou a estagnada na última década, influenciada pelas discrepâncias convergentes de tendências entre Norte/Sul que tendem a aumentar no futuro. As intervenções em estilo de vida mostraram diminuição significativa na pressão arterial de populações de países em desenvolvimento, apresentando-se como adjuvante viável ao controle deste fator de risco modificável. Entretanto, novos estudos com maior poder de análise e rigor científico são necessários.

## ABSTRACT

**BACKGROUND:** Cardiovascular diseases are the first killer in developing nations and Brazil is not an exception. Additionally, the increasing prevalence of modifiable risk factors (e.g. raised blood pressure) responds for the burden of cardiovascular diseases locally and globally.

**METHODS:** To analyze mortality trends, socioeconomic and mortality data were obtained from Instituto Brasileiro de Geografia e Estatística and from the Ministry of Health Mortality System (Datasus). Data were collected by gender, age and location. We analyzed the trends locally using Acute Myocardial Infarction (AMI) as cause of death and built a Poisson Regression model. From the model, we estimated mortality rates and number of deaths projected from the baseline rates and not observed during 1999 to 2009. Nationally, we considered Ischemic Heart Diseases as cause of death (CID I20 - I25). We included all adult residents (>20 years) living in five Brazilian regions, between 2000 and 2010. Age-standardized mortality rates (ASMRs) per 100,000 inhabitants, by sex, and region were calculated; employing a standard Brazilian population, and constructing multivariate regression models to quantify and to project temporal trends. From the adjusted rates obtained we projected future mortality trends up to 2015. For the systematic review and meta-analysis of available clinical trials, based in the Low and Middle Income Countries (LMIC) we sought for studies published before June 2012, which assessed effects of any lifestyle related intervention on blood pressure. We identified relevant studies by systematically searching multiple electronic databases. Interventions were grouped into behavioral counseling, dietary modification, physical activity and multiple interventions. We calculated and pooled effect estimates (weighted mean difference or WMD in mm Hg between the post- and pre-intervention measurements) with random-effects models. Where appropriate, subgroup and meta-regression analyses were employed to evaluate origins of heterogeneity.

**RESULTS:** In Curitiba, we found a significant trend of decline ( $p < 0.001$ ) in the period. The estimated average reduction in the rate of death from myocardial infarction each year was 3,8% (IC95%: 3,2% - 4,5%). No difference on the tendency ( $p=0,238$ ) was found between genders. In Brazil, 560,710 men and 406,627 women died due to IHD in Brazil. ASMR trends across all regions for men and women tended to converge, driven by a declining trend in the South and Southeast, and an opposite incline in the North and Northeast ( $p < 0.05$ ). Additionally, future projections demonstrated potentially widening of the observed North-South gap in coming years. The systematic review yielded 6211 references on life-style related interventions effect on blood pressure of LMIC populations. From which, 43 were included in the meta-analysis comprising 6779 non-overlapping participants. The changes (and corresponding 95% CI) in the WMD expressed as mm Hg achieved for systolic blood pressure were -5.39 (-10.73, -0.05) for behavioral counseling, -3.48 (-5.45, -1.50) for dietary modification, -11.37 (-16.06, -6.68) for physical activity and -6.09 (-8.87, -3.32) for multiple lifestyle interventions. Findings were generally consistent for diastolic blood pressure. Nonetheless, heterogeneity was generally high across these studies, which was partly explained when studies were grouped according to sample size, duration of follow-up and analytical approach used.

**CONCLUSION:** AMI mortality trend has declined in the last decade in Curitiba. In Brazil, previously reported declining IHD mortality trend changed to a stagnated trend heavily influenced by North/South discrepancies which is likely to widen further in the near future. Lifestyle interventions showed significant lowering effects on blood pressure levels of populations living in developing countries such as Brazil. Therefore they pose as an effective adjuvant therapy to control this important modifiable risk factor. Nonetheless, further investigations with sufficient power and scientific rigor would be required to more reliably quantify these effects.

## INTRODUÇÃO

A Doença Cardiovascular (DCV) é a principal causa de óbito no mundo[1]. Países desenvolvidos e países em desenvolvimento apresentam DCVs como principal causa proporcional de morte, porém 80% do impacto mundial da doença vem de países em desenvolvimento, como o Brasil [2, 3]. Nestas sociedades, a DCV representa uma ameaça ao desenvolvimento social e econômico, sobretudo devido à grande proporção de óbitos que ocorrem prematuramente [1].

Dentre as DCV, as doenças cerebrovasculares (DCbV) e as doenças isquêmicas cardíacas (DIC) são agrupadas de acordo com a morfoanatomia da patologia estabelecida. Apesar das diferenças anatômicas e etiológicas, alguns componentes causais são semelhantes entre DIC e DCbV [4, 5].

Fenômenos como o envelhecimento populacional, o tabagismo, a dieta inadequada e inatividade física estão associados à crescente prevalência de fatores de risco comuns às DIC e às DCbV como hipertensão, dislipidemias e diabetes [4, 5].

Embora desproporcionalmente distribuídos nas populações mais pobres[6] tais fatores de risco, com exceção do envelhecimento, são eminentemente evitáveis[5]. Entretanto, a reprodução de intervenções eficazes em países desenvolvidos para contextos diversos sócio economicamente como os de países em desenvolvimento não tem sido eficaz no combate às DCV [3]. Esta necessidade premente de soluções de enfrentamento das DCV adaptadas às fragilidades econômicas e estruturais das nações em desenvolvimento, tem direcionado as ações de órgãos internacionais, governos e periódicos científicos [7-9]. Por outro lado, a reunião de evidência sumarizada com precisão e poder de análise tem sido recurso cada vez mais importante no setor de saúde[10]. Revisões sistemáticas e meta análises tem sido ferramenta importante para a prevenção e tratamento das DCV[11].

No Brasil, as taxas de mortalidade por DCV apresentaram uma elevação que acompanhou a industrialização no país a partir da década de 1930 e subsequente queda a partir da década de 80 [12, 13]. Em Curitiba, um estudo de tendência de mortalidade por DIC, também evidenciou declínio a partir da década de 80[14]. No entanto, fenômeno comum observado globalmente, o aumento da prevalência de fatores de risco para DCV com incidência em indivíduos cada vez mais jovens [15,

16] ameaça as tendências de declínio, evidenciando a importância de vigilância para o planejamento adequado de ações de proteção à saúde com base em evidência adequada.

Para tanto, as tendências de mortalidade por DIC em Curitiba e no Brasil na última década não são totalmente esclarecidas. Adicionalmente, não se conhece uma revisão sistemática e/ou meta análise sobre o efeito de intervenções em estilo de vida sobre a pressão arterial de populações específicas de países em desenvolvimento.

Nesse sentido, a epidemiologia, ciência originada das observações de Hipócrates sobre fatores ambientais influenciando a ocorrência de doenças [17], apresenta-se como ferramenta para o estudo da distribuição e determinantes da doença com vistas à prevenção e ao controle dos problemas de saúde na população [18].

Portanto, tem-se com esta tese, o objetivo de analisar a DCV em nível local e nacional nos últimos 10 anos por variáveis como sexo, faixa etária e diferenças regionais, além de apresentar evidência para a prevenção e tratamento do mais importante fator de risco populacional para as DCV ,a hipertensão.

No primeiro artigo descreve-se a tendência de mortalidade por DIC em Curitiba-PR por gêneros e faixas etárias na última década. No segundo artigo apresenta-se uma análise comparativa regional das séries temporais de mortalidade por DIC por gênero e faixas etárias na última década, no contexto das disparidades socioeconômicas brasileiras e com subsequente projeção de tendências para os próximos anos. No terceiro artigo, é apresentada uma revisão sistemática e meta-análise sobre os efeitos de intervenções em estilo de vida na pressão arterial de populações de países em desenvolvimento.

## DESENVOLVIMENTO

Os artigos apresentados nesta tese foram elaborados em duas etapas. A primeira ainda no Brasil, com a coleta de dados do Datasus (Ministério da Saúde), análise e elaboração do manuscrito sobre mortalidade por IAM em Curitiba em parceria com o Centro de Epidemiologia da Secretaria Municipal da Saúde de Curitiba representado pela Prof. Dra. Karin Regina Luhm. Este trabalho resultou em um artigo publicado em periódico nacional e uma apresentação oral em congresso regional.

Como resultado desta análise, nosso grupo ganhou experiência com a coleta e análise de dados de mortalidade utilizando as ferramentas do Datasus (Tabnet e Tabwin). A análise da mortalidade por outra importante causa de óbito em nosso meio, a doença cerebrovascular, resultou em dois resumos apresentados em congresso e evidenciou diferenças em termos de qualidade de informação e comportamento destas tendências na cidade de Curitiba, no mesmo período. O próximo passo, consistiu na comparação das tendências de mortalidade por DIC e DCbV em Curitiba, na última década, evidenciando que diferente de outras capitais do país, Curitiba ainda apresenta a DIC como primeira causa de óbito, entretanto com declínio paralelo entre as duas causas de óbito (DIC e DCbV) para os dois gêneros e grupos etários, exceto para mulheres acima de 60 anos. A tendência de declínio neste grupo, foi mais acentuada por DIC. Esta última análise foi apresentada como resumo em um congresso nacional e um congresso internacional.

Nesta fase, durante o programa de verão em epidemiologia na Erasmus MC em Rotterdam- Holanda, o Prof. Dr. José Rocha Faria Neto delimitou com o Professor Dr Oscar Franco, daquela instituição, um período de estágio doutoral para pesquisa no exterior.

Ainda neste período, foram coletados dados de mortalidade por DIC e DCbV nas diversas regiões brasileiras durante a última década e estas tendências foram testadas para correlação com dados socioeconômicos do mesmo período nas respectivas regiões. O resultado mostrou que o aumento de renda observado em todo o país na última década apresentou relações lineares inversas com as tendências de óbito nas regiões mais ricas enquanto nas regiões mais pobres, esta relação foi paralela. Estes achados resultaram em um resumo apresentado em um congresso mundial e uma apresentação oral em congresso nacional.

Justificou-se assim, a necessidade de uma análise aprofundada das tendências regionais de mortalidade por Doença Isquêmica Cardíaca no contexto sócio econômico que seria realizada em colaboração com duas universidades estrangeiras.

O início do estágio no Departamento de Epidemiologia, Grupo de Doenças Cardiovasculares e Envelhecimento Saudável da Erasmus MC, marcou a segunda etapa desta tese. Os objetivos a serem atingidos neste período seriam um artigo sobre a mortalidade por Doença Isquêmica Cardíaca nas regiões brasileiras e uma meta análise sobre a eficácia de intervenções não farmacológicas na prevenção e tratamento da pressão arterial elevada em países em desenvolvimento. Para isso, foram coletados os dados sócio demográficos e de mortalidade por região, sexo e faixa etária. A análise e a elaboração do manuscrito foram realizadas com a colaboração do Prof. Dr. Oscar Franco da Universidade Erasmus MC e do pesquisador associado da Universidade de Cambridge Dr. Rajiv Chowdhury.

Os dois projetos foram conduzidos durante o estágio no exterior. Para a construção do conhecimento necessário para os dois projetos, além das disciplinas cursadas no Brasil, foram cursados workshops sobre meta análise, curso de verão em Meta análise e Análise de Regressão, workshops sobre uso de gerenciador de bibliografia (Endnote), pesquisas bibliográficas em Pubmed e outras bases de dados (Embase, Ovid SP, Cochrane, Web of Science, CINAHL e Scopus).

Durante a análise dos dois projetos, foi ainda realizado um estágio na Universidade de Cambridge, Escola de Saúde Pública para treinamento em ferramentas de análise e apresentação gráfica de resultados obtidos em diferentes programas estatísticos, análise de subgrupos e metaregressão.

Como resultado desta colaboração, os dados de mortalidade por DIC no Brasil foram apresentados de forma diferente dos trabalhos até então realizados na área e foi evidenciada a grande disparidade regional, sua influência na tendência do país além de projeções para o futuro que justificam a necessidade de pesquisas desenhadas regionalmente para o combate desta importante causa de morbimortalidade em nosso meio.

Diante dos resultados obtidos com a análise do Brasil, identificou-se o efeito de fatores socioeconômicos na apresentação da doença cardiovascular em um país em reconhecida ascensão econômica. A escassez de evidência focada em países em desenvolvimento motivou a condução da revisão sistemática e meta análise que se seguiu. Apesar das evidências sobre custo eficácia de tratamentos farmacológicos para o

mais prevalente fator de risco para DCV em países em desenvolvimento, a hipertensão; este segue sendo um dos grandes desafios nos setores primário e secundário da saúde em nível populacional.

Diante da inexistência de evidência reunida e sumarizada sobre terapias adjuvantes para a prevenção e tratamento desta doença/ fator de risco nesta população específica, optou-se por analisar ensaios clínicos de intervenções não farmacológicas conduzidas exclusivamente em países em desenvolvimento.

O processo de construção da estratégia de pesquisa e da recuperação de artigos originais envolveu um bibliotecário especialista em bibliografia médica, Wichor Brammer que auxiliou na ampliação da busca para bibliotecas remotas e bases de dados menos comumente usadas para revisões sistemáticas, mas que no caso de evidência focada em países em desenvolvimento, seriam indispensáveis. Um número relativamente grande de referências foi reunido (aproximadamente 9000) e para tanto, mais 12 revisores de diversas nacionalidades foram convidados ao grupo de trabalho.

A revisão sistemática resultou em um número de intervenções comparável a outras revisões sobre o tema, evidenciando que há um grande número de descrições de intervenções em estilo de vida originadas de nações em desenvolvimento, porém a meta análise, sobretudo as análises de subgrupo, evidenciaram a necessidade premente do uso de guias e padronização na descrição destes ensaios clínicos.

No decorrer das duas etapas deste conjunto, a análise da mortalidade por DCV em nível local subsidiou a análise em nível nacional que por sua vez, ressaltou a influência de fatores socioeconômicos na mortalidade por DCV e motivou a busca de evidência em medidas alternativas de combate ao grande impacto que este importante fator de risco, a pressão arterial elevada exerce em populações vulneráveis, das quais o Brasil não é exceção.

Serão apresentados a seguir, aspectos metodológicos relevantes dos três artigos, seguidos pelos manuscritos.

## **METODOLOGIA**

Neste conjunto de artigos, diferentes métodos de epidemiologia básica e epidemiologia clínica foram utilizados para a análise da mortalidade por DIC em nível local e nacional, além da reunião de evidência sobre controle de fator de risco cardiovascular. No primeiro artigo, conduziu-se um estudo observacional de tendência temporal de mortalidade. No segundo artigo, foi realizado um estudo populacional comparativo de delineamento ecológico. No terceiro artigo, utilizou-se uma revisão sistemática e a técnica de meta-análise para a apresentação de evidência reunida sobre o controle da hipertensão em países em desenvolvimento. A descrição de cada metodologia segue.

### **a) Artigo I**

No primeiro artigo, foram obtidos do Sistema de Informação de Mortalidade (SIM) do Departamento de Informática do Sistema Único de Saúde (DATASUS)/Ministério da Saúde (MS). Para a tabulação dos dados, utilizou-se o Sistema de Tabulação do DATASUS para Windows, *Tabwin*.

A pergunta que definiu as variáveis a serem extraídas do Sistema de Informação de Mortalidade foi “Qual é o número de óbitos por IAM ocorridos em Curitiba entre homens e mulheres adultos em cada ano da última década?”. Na sequência, buscou-se o arquivo de definição (.DEF) e os arquivos de base de dados (.DBF) para óbitos em Curitiba em cada um dos anos estudados. Para as opções de tabulação, utilizou-se Ano do óbito na opção linha, Sexo na opção coluna e Faixa etária na opção Incremento. Na opção Seleções disponíveis, utilizou-se Curitiba.

Para a extração de dados sobre causa de mortalidade considerou-se a Causa CID BR-10 com código 068.1 equivalente ao código I-21 do Código Internacional de Doenças, CID-10. As faixas etárias analisadas foram de 20 a 49 anos, 50 a 59 anos, 60 a 69 anos, 70 a 79 anos e 80 anos ou mais. Os dados de óbito foram coletados por local de residência.

Os dados demográficos foram obtidos do Instituto Brasileiro de Geografia e Estatística (IBGE), sendo os denominadores dos cálculos de taxa correspondentes à população por gênero e faixa etária de acordo com os dados do período de 1998 a 2009.



Os dados foram extraídos no formato de planilha Excel e traduzidos para o formato específico do Pacote de análise estatística SPSS.

Construiu-se um modelo de Regressão de Poisson considerando-se como variável resposta o número de óbitos e como variável explicativa o tempo correspondente aos anos observados. A partir do ajuste do modelo de Poisson e tendo como base o primeiro ano do estudo (1998) foi estimado o número de mortes por IAM que seriam esperadas para o período de 1999-2009, mas que não ocorreram.

## **b) Artigo II**

No segundo artigo, a população do estudo envolveu todos os adultos (acima de 20 anos) residentes nas cinco regiões brasileiras, por sexo, com extratos de dez anos para faixas etárias, no período de 2000 a 2010. O desfecho estudado foi o número de óbitos por DIC com causas de óbitos codificadas como I-20 a I25 pelo Código Internacional de Doenças CID-10, correspondentes a angina pectoris, infarto agudo do miocárdio (IAM), subseqüentes complicações de IAM, outras doenças isquêmicas cardíacas agudas e doenças isquêmicas cardíacas crônicas. Os óbitos por causas não esclarecidas foram coletados para cada ano, região, sexo e faixa etária e a estes, foram aplicados a proporção de óbitos pelas causas de interesse. Na seqüência, foi realizada a correção dos óbitos por IAM para causas mal definidas. As taxas de mortalidade específicas por idade foram calculadas para cada ano, gênero e região tendo no numerador o número de óbitos em cada faixa etária, gênero, ano e região e no denominador, a população correspondente.

A padronização tendo como padrão a população brasileira foi calculada aplicando-se a proporção da população brasileira no correspondente ano, gênero e faixa etária à taxas específicas por idade e gênero de cada região. Na seqüência, somou-se as novas taxas obtidas para todas as faixas etárias (20-80 anos ou mais) e multiplicou-se por 100.000, obtendo-se assim a Taxa de mortalidade padronizada por idade de cada região, e em cada ano adaptada para comparação entre populações que apresentam diferentes distribuições etárias.

Os dados demográficos e os dados de renda média per capita foram obtidos do Instituto Brasileiro de Geografia e Estatística (IBGE). Para a comparação de tendências entre regiões, construiu-se um modelo hierárquico de regressão linear multivariada. Foram codificadas variáveis “dummy” para regiões e as tendências de óbito foram

testadas para diferenças quanto a regiões, sexos e faixas etárias. A mudança anual média nas tendências de mortalidade por sexos e faixa etária foi calculada a partir de um modelo de regressão de Poisson tendo como variável dependente a taxa de óbito específica por idade, o tamanho populacional como variável de peso e cada ano do período foi utilizado como variável independente. Subsequentes projeções das taxas de mortalidade padronizadas por idade para os anos 2011, 2012, 2013, 2014 e 2015 foram calculados a partir do ajuste do modelo hierárquico por região e sexo.

### **c) Artigo III**

No terceiro artigo conduziu-se uma revisão sistemática utilizando-se uma combinação dos termos Medical Subject Headings (MeSH) e palavras chave que incluíram a população (ex. *developing country* ou *low and middle income country*), termos relacionados a intervenções em estilos de vida (ex. diet, physical activity, education) e termos relacionados ao desfecho de interesse (ex. *blood pressure, hypertension*). Os termos de busca foram extraídos de outras revisões sobre assuntos semelhantes, e estudos relevantes nas áreas de interesse. Na sequência, buscou-se cada um dos termos e sinônimos nas definições de Medical Subject Headings (MeSH) e designados na estratégia de busca como referentes às categorias População, Intervenção, Desfecho e Comparações analisadas.

A seguir, designou-se a cada termo o nível de importância nos estudos pesquisados como título e resumo ou em qualquer lugar do texto. Entre termos sinônimos, acrescentou-se a preposição OU(OR), entre termos não sinônimos, acrescentou-se a preposição E(AND) e entre grupos de termos referentes à população, intervenções e desfecho, acrescentou-se a preposição E(AND). A construção da estratégia de busca levou aproximadamente 300 horas de trabalho.

Na sequência, a estratégia de busca originada na base de dados PUBMED foi traduzida para as demais bases de dados conforme as diferentes características de busca de cada uma delas. As referências foram extraídas das seguintes bases de dados: Medline-Pubmed, Embase, Cochrane Library, CINAHL, Web of Science, Scopus, Scielo e LILACS. Adicionalmente, foram obtidas referências de outras revisões sistemáticas. Foram extraídas as referências publicadas entre janeiro, 1977 a setembro, 2012. Inicialmente, foram obtidas aproximadamente 9000 referências, e após de duplicação de referências repetidas em diferentes bases de dados, resultaram em 6211 referências.

Os títulos e resumos das 6211 referências foram revisadas em duplas aplicando os critérios de seleção previamente definidos e desacordos foram resolvidos por um terceiro revisor independente. No total, foram envolvidos 12 revisores neste processo. O processo de revisão dos artigos tomou aproximadamente 400 horas.

Informações sobre todas as possíveis causas de heterogeneidade entre os estudos incluídos foram registradas em um formulário de coleta de dados e transcritas para uma planilha de dados do tipo Excel. Nesta planilha, cada artigo representava um linha e cada coluna representava uma variável de interesse para análise descritiva ou inferencial. No caso de ausência de informações relevantes para a análise, autores foram contatados duas vezes. Após segundo contato sem resposta, a informação era dada como ausente. A preparação do banco de dados tomou aproximadamente 250 horas.

As intervenções foram agrupadas em intervenções educacionais, atividade física, modificações de dieta e intervenções combinadas. A diferença média da pressão arterial foi calculada subtraindo-se a pressão arterial média pós-intervenção da pressão arterial média pré-intervenção com respectiva variância (IC 95%) no grupo controle e no grupo intervenção de cada estudo. Na sequência, subtraiu-se a diferença média do grupo controle da diferença média do grupo intervenção de cada estudo com respectiva variância (IC 95 %), este resultado é o chamado efeito da intervenção. Foi realizada uma meta análise pelo método do modelo “*random*” com os dados extraídos dos artigos. Foram realizadas análise de subgrupos de acordo com possíveis causas de heterogeneidade como tamanho da amostra, qualidade dos estudos, uso de drogas anti-hipertensivas, presença de co-morbidades, duração da intervenção, etnias e características das intervenções. Também foi utilizada meta regressão para explicar a heterogeneidade entre os estudos para variáveis contínuas como idade média e proporção de mulheres. As análises e gráficos foram gerados no pacote estatístico Stata versão 12.0 (StataCorp LP, College Station, Tex). A análise tomou aproximadamente 120 horas.

A seguir serão apresentados os três artigos mencionados.

## **ARTIGO 1**

**Tendência de mortalidade por Infarto Agudo do Miocárdio em Curitiba-PR no período de 1998 a 2009**

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## **Resumo**

**Fundamento:** O Infarto Agudo do Miocárdio é a principal causa isolada de óbito entre as doenças crônicas não transmissíveis no Brasil. O conhecimento das tendências de mortalidade é necessário para o planejamento de estratégias de prevenção.

**Objetivo:** Avaliar a tendência de mortalidade por infarto do miocárdio no período de 1998 a 2009 na cidade de Curitiba, sua distribuição por gênero, faixa etária e seu impacto na redução do número absoluto de mortes por esta patologia nesse período.

**Métodos:** Dados demográficos foram obtidos do Instituto Brasileiro de Geografia e Estatística e dados de óbitos foram obtidos no Sistema de Informação de Mortalidade do Ministério da Saúde considerando gênero, faixa etária e residência. A partir do ajuste de um modelo de Regressão de Poisson foram estimadas taxas de mortalidade e número de mortes esperadas que não foram observadas.

**Resultados:** Foi encontrada tendência de declínio significativa ( $p < 0,001$ ) no período. A estimativa da redução média na taxa de óbito por IAM a cada ano foi de 3,8% (IC95%: 3,2% - 4,5%). Não houve diferença significativa entre os gêneros ( $p = 0,238$ ), entretanto, a evolução das taxas padronizadas específicas por idade diferiu significativamente entre as faixas etárias ( $p = 0,018$ ). Estima-se que a redução anual de 3,8% na taxa de mortalidade tenha resultado em 2.168 mortes aquém do número esperado, considerando a taxa de mortalidade observada em 1998 e projetando este número sobre o crescimento populacional ocorrido no período estudado.

**Conclusões:** Embora permaneça como causa importante de óbito, a mortalidade por IAM apresentou queda significativa no período avaliado

## **Abstract**

**Background:** Acute Myocardial Infarction is the leading single cause of death among chronic non-communicable diseases in Brazil. The knowledge of mortality trends is necessary for planning preventive strategies.

**Objectives:** To evaluate the mortality rate trends of myocardial infarction from 1998 to 2009 in Curitiba, and its distribution by gender, age and also the impact on the number of deaths in this period.

**Methods:** Demographic data were obtained from the Brazilian Institute of Geography and Statistics. Mortality data were obtained from the Mortality Information System of the Health Ministry by gender, age and residence. A Poisson regression model was adjusted. Rates were calculated for the evolution of mortality models that adjusted to Poisson distribution. From this model, the number of deaths prevented or postponed was calculated.

**Results:** A significant trend of decline ( $p < 0.001$ ) was found in the period. The estimated average reduction in the rate of death from myocardial infarction each year was 3,8% (IC95%: 3,2% - 4,5%). No difference on the tendency ( $p=0,238$ ) was found between genders, however, the evolution of age-specific mortality rates differed significantly ( $p=0,018$ ) between age groups. An estimated 2.168 deaths less than expected from the baseline occurred during this time .

**Conclusions:** Although it remains as the major single cause of death, mortality from AMI decreased significantly in the evaluated period. There is a need for a detailed analysis of the factors associated with this decline.

## INTRODUÇÃO

As doenças cardiovasculares (DCV) permanecem como principal causa de morte nos países desenvolvidos e nos em desenvolvimento<sup>1</sup>, embora se observe nas últimas décadas um declínio nesta taxa de mortalidade<sup>2,3</sup>. Há, entretanto, evidências de diferenças importantes nesta queda em relação à distribuição geográfica, faixa etária, gênero, etnias e nível sócio-econômico<sup>4,5</sup>.

No Brasil, as taxas de mortalidade por DCV apresentaram uma elevação que acompanhou a industrialização no país a partir da década de 1930. Dentro do grande grupo de DCV, as Doenças Isquêmicas do Coração (DIC) são as causas de óbito mais ocorrentes, sendo o infarto agudo do miocárdio (IAM) a causa isolada de morte mais comum em homens e mulheres<sup>6</sup>. Entretanto, observa-se queda do risco de óbito por DCV ajustado por idade a partir da década de 90 nas regiões Sul, Sudeste e Centro-Oeste e algumas capitais do Norte e Nordeste, com algumas diferenças entre gêneros<sup>7</sup>.

A manutenção de tendência de declínio, no entanto, parece ser questionável, uma vez que a prevalência de alguns fatores de risco, como obesidade e Diabetes Mellitus, tem aumentado<sup>8,9</sup>. Estudos com grandes populações demonstram que o risco atribuído a estes fatores é significativo<sup>10,11</sup>. Por outro lado, a análise do impacto exercido pelo controle de outros fatores de risco, seja em prevenção primária ou secundária, demonstra que o controle destes fatores é determinante para a redução da mortalidade cardiovascular que tem sido observada<sup>3, 12 13</sup>. Nos Estados Unidos, metade da redução da mortalidade cardiovascular em duas décadas pôde ser explicada por um melhor controle dos fatores de risco, enquanto a outra metade foi atribuída aos tratamentos específicos das patologias específicas<sup>14</sup>.

No município de Curitiba-PR, o Infarto Agudo do Miocárdio tem sido a principal causa de morte isolada nos últimos 10 anos<sup>15</sup>, porém a taxa de mortalidade ajustada por idade e gênero no mesmo período não é descrita. Estudo prévio sobre doenças isquêmicas do coração nesta localidade durante o período de 1980 a 1998, evidenciou diferenças significativas entre os gêneros e idades<sup>16</sup> em relação à mortalidade por IAM. Neste sentido, a avaliação local da evolução na tendência de mortalidade nos anos subsequentes é fundamental para o planejamento de políticas públicas de saúde e planejamento de ações de promoção e prevenção a serem executadas por entidades públicas e privadas.

O objetivo deste estudo foi avaliar a evolução da taxa de mortalidade por infarto agudo do miocárdio no período de 1998 a 2009 na cidade de Curitiba, bem como a distribuição da mortalidade por gênero, faixa etária e seu impacto no número absoluto de mortes por IAM neste período.

## **MÉTODOS**

Neste estudo do tipo observacional ecológico, os dados sobre causas de mortes no período de 1998 a 2009 foram obtidos do Sistema de Informação de Mortalidade (SIM) do Departamento de Informática do Sistema Único de Saúde (DATASUS)/Ministério da Saúde (MS).

Para a extração de dados sobre causa de mortalidade considerou-se a Causa CID BR-10 com código 068.1 equivalente ao código I-21 da CID-10<sup>17</sup>. As faixas etárias analisadas foram de 20 a 49 anos, 50 a 59 anos, 60 a 69 anos, 70 a 79 anos e 80 anos ou mais. Os dados de óbito foram coletados por local de residência.<sup>6</sup>

Os dados demográficos foram obtidos do Instituto Brasileiro de Geografia e Estatística (IBGE)<sup>18</sup>, sendo os denominadores dos cálculos de taxa correspondentes à população por gênero e faixa etária de acordo com os dados do período de 1998 a 2009.

### **Análise Estatística**

Para avaliação das taxas de mortalidade, ajustou-se um modelo de Regressão de Poisson considerando-se como variável resposta o número de óbitos e como variável explicativa o tempo correspondente aos anos observados. Como variável de exposição foi considerada a população em cada ano avaliado. A função de ligação foi a exponencial e, para avaliação do ajuste, foi considerada a função *deviance*.

O teste de Wald foi usado para avaliação da importância do efeito do tempo sobre a taxa de óbito. Este mesmo teste foi considerado para avaliação do paralelismo entre grupos em relação à variação da taxa de mortalidade ao longo do tempo. Quando da identificação desta importância, estimou-se a taxa de variação média dos anos consecutivos pelo modelo, com o respectivo intervalo de 95% de confiança.

A partir do ajuste do modelo de Poisson e considerando-se a linha de base (1998) foi estimado o número de mortes por IAM que seriam esperadas para o período de 1999-2009, mas que não ocorreram.



Valores de  $p < 0,05$  indicaram significância estatística. A análise foi realizada com o programa computacional SPSS v.14.0.

## **RESULTADOS**

O Infarto Agudo do Miocárdio permanece como primeira causa isolada de óbito entre as doenças crônicas não transmissíveis apresentando mortalidade proporcional de 9,1% em 1998 e 6,7% em 2009. Porém, ressalta-se que a partir de 2003, o IAM não configura a primeira causa isolada, tendo sido superada pelas mortes por causas externas.

Os resultados do modelo geral de mortalidade por IAM no período de 1998 a 2009 indicaram declínio significativo na taxa de mortalidade por IAM no período avaliado ( $p < 0,001$ ) e a estimativa da redução média nessa taxa a cada ano foi de 3,8% (IC95%: 3,2% - 4,5%).

Na Figura 1 são apresentadas as taxas de mortalidade/100.000 indivíduos observados e as taxas estimadas pelo ajuste do modelo de Poisson com respectivos intervalos de 95% de confiança.

### **Taxa de óbito por IAM: análise por gênero**

Ao longo do período estudado, a razão entre a taxa de óbito do gênero masculino e a taxa para o gênero feminino é 1,46.

Para o gênero masculino, evidenciou-se uma tendência de declínio ( $p < 0,001$ ) com estimativa da redução média na taxa de óbito por IAM a cada ano de 3,5% (IC95%: 2,7% - 4,3%). Da mesma maneira, para o gênero feminino houve tendência de declínio ( $p < 0,001$ ) com estimativa da redução média na taxa de óbito por IAM a cada ano de 4,2% (IC95%: 3,3% - 5,2%).

Adicionalmente, testou-se o paralelismo entre o gênero masculino e feminino em relação à evolução das taxas de mortalidade. Os resultados encontrados indicaram que não houve diferença significativa ( $p = 0,238$ ) entre os gêneros em relação à tendência, conforme demonstra a **Figura 2**.

### **Taxas padronizadas específicas por idade**

As faixas etárias de 20 a 29, 30 a 39 e 40 a 49 anos foram agrupadas em uma nova categoria de 20- 49 anos devido ao pequeno número de ocorrências em algumas

das faixas etárias citadas. Nesta categoria, encontrou-se decréscimo significativo ( $p < 0,001$ ), com redução média na taxa de óbito por IAM de 7,4% (IC95%: 5,2% - 9,6%) por ano. Na faixa etária de 50 a 59 anos o declínio também foi significativo ( $p < 0,001$ ), com estimativa da redução média na taxa de óbito por IAM por ano de 7,0% (IC95%: 5,5% - 8,4%).

Também nas faixas etárias mais elevadas, 60-69, 70-79 e 80 anos ou mais o decréscimo foi significativo ( $p < 0,001$ ). A estimativa da redução média na taxa de óbito por IAM a cada ano foi de 6,6% (IC95%: 5,4%-7,8%) nos indivíduos de 60-69 anos, de 7,2% (IC95%: 6,1% - 8,4%) nos indivíduos de 70-79 anos e de 4,3% (IC95%: 3,0% - 5,5%) nos indivíduos de 80 anos ou mais.

A comparação entre as faixas etárias de 20-49, 50-59, 60-69, 70-79 anos e 80 anos ou mais em relação ao declínio na taxa de mortalidade foi realizada testando-se a hipótese de paralelismo. A evolução da taxa de mortalidade na faixa de 80 anos ou mais difere significativamente dessa evolução para as demais faixas etárias ( $p = 0,018$  para 20-49 anos;  $p = 0,008$  para 50-59 anos;  $p = 0,012$  para 60-69 anos;  $p = 0,002$  para 70-79 anos). Entretanto, nas demais comparações entre as faixas etárias, não foi encontrada diferença significativa quanto à evolução das taxas de mortalidade por IAM (Figura 3).

### **Número de mortes a menos do que o esperado a partir da linha base de 1998**

A partir do ajuste do modelo de Poisson, o número estimado de mortes por IAM no período de 1998-2009, considerando-se as taxas ajustadas de cada ano, é de 9.065. Entretanto, se a taxa ajustada de 1998 fosse mantida, o número de mortes estimado seria de 11.233. Esses resultados indicam que, mantendo-se o declínio médio de 3,8% no período de 12 anos, estima-se que 2.168 mortes seriam esperadas, mas não foram observadas nesse período (Tabela 1). A evolução do número acumulado de mortes esperadas e não observadas no período analisado é apresentada na Figura 4. O declínio da mortalidade por IAM pode ainda ser representado pela diminuição do risco de morte usado nas tábuas de vida. Neste estudo, o risco de morte observada por IAM em Curitiba em 2009 foi 38,2% mais baixo do que em 1998.

### **DISCUSSÃO**

O estudo da mortalidade é utilizado como medida de parâmetros de saúde da população, e o delineamento de estudo ecológico tem como característica, a determinação geográfica da população estudada <sup>19</sup>. Sabe-se que este tipo de estudo não se propõe à análise em nível individual e nem ao estabelecimento de relações causais.

Neste estudo, uma possível causa de viés, a idade, foi controlada através da padronização específica. O que se apresenta portanto, é a linha temporal de uma causa de mortalidade da população de Curitiba, que não deve ser inferida a outras populações, porém pode ser comparada com outras populações e pode ainda, basear estudos longitudinais de relações causais.

A mortalidade por infarto agudo do miocárdio apresentou significativa redução na cidade de Curitiba no período avaliado. A redução foi consistente ao longo dos anos, em ambos os gêneros e em todas as faixas etárias abaixo dos 80 anos. Esta redução resultou, ao término do período avaliado, em 2.168 mortes a menos do que poderia se esperar projetando a taxa de mortalidade de 1998 e levando-se em consideração o crescimento populacional no mesmo período. Ressalta-se que a redução ocorreu a despeito do aumento das taxas de admissões hospitalares por IAM no período, achado este que corroborado por outros autores nos últimos 20 anos<sup>5</sup>. A população com mais de

20 anos residente no município de Curitiba cresceu 19,5% no período estudado (1998 a 2009) e as taxas de hospitalizações por IAM pelo SUS aumentaram 35%.

O declínio anual das taxas de mortalidade por doenças isquêmicas do coração tem sido descrito em capitais do Brasil<sup>20,21</sup>. Estudo anterior que analisava as tendências de mortalidade por IAM e Doenças Isquêmicas do Coração em Curitiba entre 1980 e 1998<sup>16</sup> já evidenciava tendência de diminuição de mortalidade por IAM, porém em ritmo de decréscimo menor do que os 3,8 % ao ano aqui demonstrados. Embora aquele estudo tenha utilizado outro método de análise para a tendência, nossas estimativas de percentual anual de decréscimo foram realizadas em relação aos anos imediatamente anteriores evidenciando uma tendência de declínio ainda maior do que a encontrada em período anterior. Outro aspecto a ser considerado é o da limitação das projeções populacionais intercensitárias. Nosso estudo utilizou os dados do Censo de 2010 que corrigiu as projeções anteriores e evidenciou a superestimação dos dados antes apresentados.

Em relação à diferença entre os gêneros, a proporção de mortes masculino/feminino, encontrada em nosso estudo foi, em média, de 1,46 enquanto naquele trabalho, a proporção encontrada foi 1,6. Outros estudos realizados em capitais brasileiras que também têm evidenciado a tendência de queda da mortalidade por IAM em período semelhante, apontam para a diferenças entre os gêneros<sup>22</sup>. Dados do estudo INTERHEART<sup>11</sup> demonstraram que mulheres tendem a sofrer o primeiro infarto mais tarde do que homens, porém este fenômeno não parece refletir-se na tendência de declínio de mortalidade. Em nosso estudo, a diminuição foi semelhante entre os gêneros, evidenciada pelo teste de paralelismo sugerindo que a maior tendência de diminuição de mortalidade masculina por IAM registrado anteriormente parece ter se direcionado a um paralelo em relação à mortalidade feminina na última década.

Na comparação entre faixas etárias, as idades mais avançadas (60-69 e 70-79 anos) apresentaram maior proporção de declínio em relação às idades mais jovens (20-49 anos) indicando que a diminuição da tendência de mortalidade por IAM observada neste estudo parece caminhar para ao quarto estágio do modelo de transição epidemiológica adaptado para doenças cardiovasculares descrito por Yusuf e cols.<sup>23</sup> Nesta fase, os esforços direcionados ao diagnóstico e ao tratamento de doenças cardiovasculares conseguem atrasar a mortalidade para idades mais avançadas. Este achado sugere uma fase de transição epidemiológica das doenças cardiovasculares para

Curitiba caso o mesmo fenômeno seja observado em outras causas de óbito cardiovascular, diferente de outras regiões metropolitanas do Brasil, e mais próxima de proporções encontradas em países desenvolvidos.<sup>24</sup>

Ainda em relação às idades, a não adequação da faixa de 80 anos ou mais nas comparações entre essa faixa e as demais à distribuição de Poisson parece ter sido efeito do comportamento errático em alguns anos do período, porém observa-se a tendência de declínio. Além da elevação prevista nas estratificações de risco, é interessante ressaltar que especialmente nesta faixa etária, os fatores socioeconômicos parecem estar mais relacionados com a maior dificuldade de declínio de mortalidade por DCV como tem sido descrito<sup>20, 25</sup>. Embora se saiba que este fenômeno possa estar influenciando o comportamento da tendência de mortalidade, esta associação não foi analisada neste trabalho. Há que se considerar ainda, que a faixa etária de 80 anos ou mais não está incluída na Lista Brasileira de Mortes Evitáveis, uma vez que a metodologia daquela lista se baseia na expectativa de vida da população sendo 75 anos, a idade limite da lista atual<sup>26</sup>.

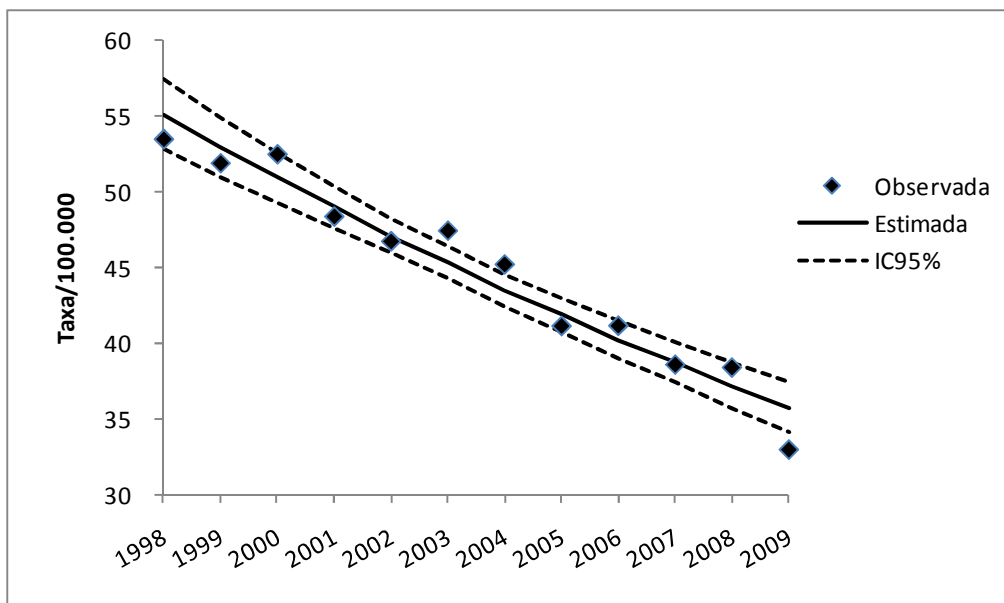
Este estudo restringiu-se à análise dos dados de óbito, e permanecem obscuras as causas do declínio da mortalidade por IAM. Parte significativa do risco de IAM e de doenças cardiovasculares está associado a fatores modificáveis amplamente conhecidos<sup>27, 28</sup>. Segundo dados do estudo INTERHEART<sup>11</sup>, globalmente pode-se atribuir 90% do risco de um primeiro IAM à presença de seis fatores de risco (dislipidemia, hipertensão, tabagismo, diabetes, obesidade abdominal e fatores psicossociais) ou ausência de 3 fatores “protetores” (consumo diário de frutas e verduras, atividade física e consumo leve de álcool). Estudos populacionais que avaliaram os fatores relacionados ao declínio da mortalidade cardiovascular, seja em prevenção primária ou secundária, demonstram que o controle destes fatores de riscos, e não só a melhora no tratamento das síndrome agudas, respondem por parcela significativa no quadro de declínio de mortalidade.<sup>12, 13, 29</sup>

Neste sentido, o cálculo do número de mortes a menos do que o esperado a partir da linha base de 1998 serve como ponto de partida para modelos de análise que avaliem o peso do controle de fatores de risco e o impacto de terapias efetivas amplamente utilizadas. Alguns estudos apontam o número de mortes evitadas a partir da terapia fibrinolítica<sup>30</sup> no manejo do IAM, porém a combinação e o peso das terapias atualmente preconizadas como trombólise, antiplaquetárias,  $\beta$  –bloqueadores, inibidores da Enzima

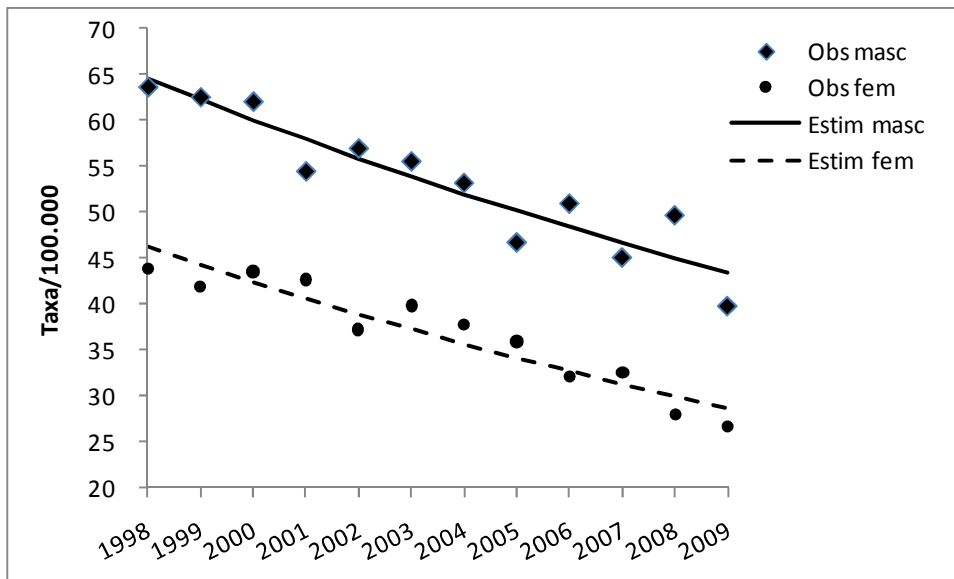
Conversora de Angiotensina e angioplastia no número de vidas salvas não é claro. Em nosso meio, estes componentes ainda não foram analisados simultaneamente. A carência de dados registrados, provenientes dos serviços públicos e privados, sobre a prevalência dos diversos fatores de risco modificáveis ao longo dos anos, coloca em dúvida a possibilidade desta análise ser realizada de maneira confiável em larga escala no nosso País.

Em conclusão, a tendência de declínio da mortalidade por IAM em Curitiba/PR no período de 1998 a 2009 foi significativa, evidenciando uma diminuição de 38,2 % no risco de morte por IAM em indivíduos de 20 a 80 anos ou mais. Esta diminuição resultou em 2.168 mortes esperadas e não observadas em no período. Uma análise detalhada dos fatores associados a esta redução seria necessária para futuros planejamentos de ações nos diferentes níveis de atenção à saúde em nosso meio.

**Figura 1-** Modelo geral da tendência de mortalidade por IAM em Curitiba-PR no período de 1998 a 2009

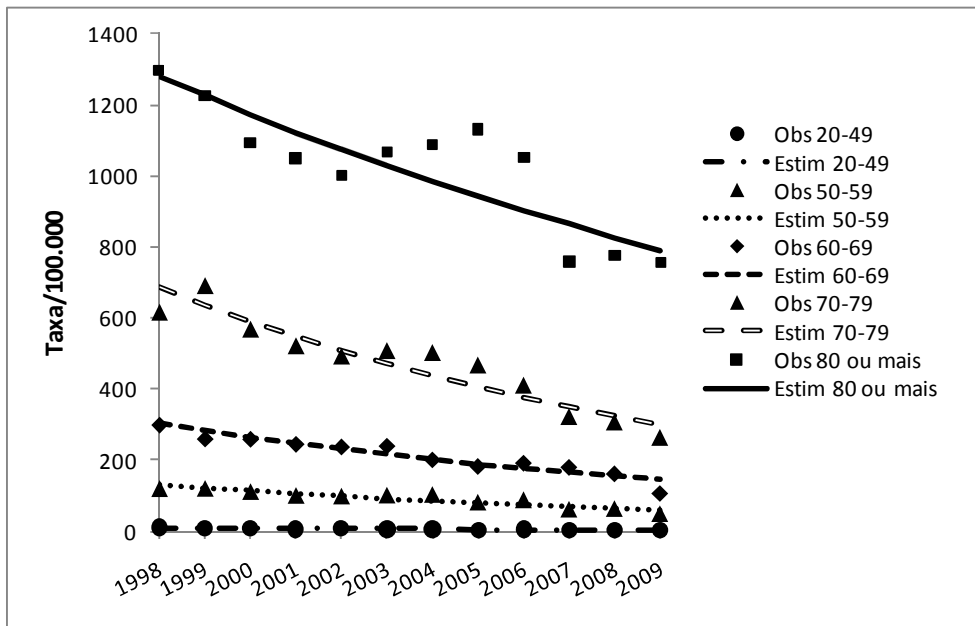


**Figura 2-** Tendência da mortalidade por IAM no período de 1998 a 2009 em Curitiba -PR para ambos os gêneros.





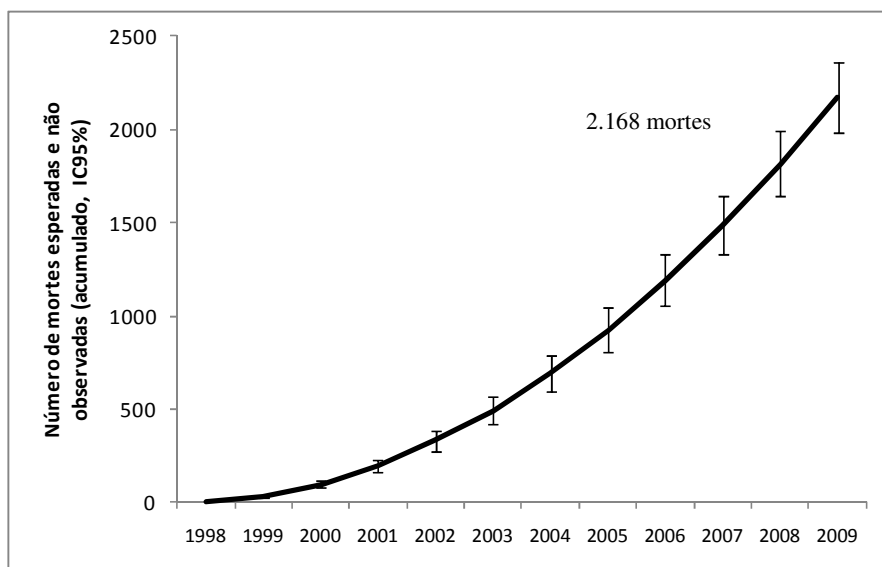
**Figura 3-** Tendência de mortalidade por IAM em Curitiba-PR no período de 1998 a 2009 para as diferentes faixas etárias



**Tabela 1** – Mortes esperadas a partir da linha base de 1998 e não observadas no período 1998 a 2009.

<b>Ano</b>	<b>População</b>	<b>Mortes estimadas</b>	<b>Mortes esperadas sem declínio (taxa de 1998)</b>	<b>Mortes esperadas e não observadas com o declínio (IC95%)</b>
1998	1.550.315	854	-	-
1999	1.584.232	839	872	34 (28 - 40)
2000	1.587.315	808	874	66 (55 - 78)
2001	1.620.221	793	892	99 (84 - 115)
2002	1.644.599	774	906	131 (113 - 151)
2003	1.671.193	756	920	164 (143 - 186)
2004	1.697.703	739	935	196 (174 - 219)
2005	1.757.903	736	968	232 (211 - 255)
2006	1.788.560	720	985	265 (245 - 286)
2007	1.818.950	704	1.002	298 (280 - 316)
2008	1.828.092	680	1.007	326 (311 - 343)
2009	1.851.213	662	1.019	357 (344 - 370)
<b>Total</b>	<b>20.400.296</b>	<b>9065</b>	<b>11.233</b>	<b>2168 (1988 – 2359)</b>

**Figura 4** – Acumulado de mortes esperadas e não observadas a partir da linha base entre 1998 e 2009 em Curitiba-PR.



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## ARTIGO 2

### **Ischemic heart disease deaths in Brazil:**

#### **Current trends, regional disparities and future projections**

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**Aceito-Heart/ BMJ**

**Ischemic heart disease deaths in Brazil:  
Current trends, regional disparities and future projections**

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ABSTRACT (word count: 245)

**Objective:** To quantify the recent trend of Ischemic Heart Diseases (IHD) deaths in Brazil during the last decade (2000-2010), to characterize these trends for various population characteristics and to forecast the upcoming mortality trends across various regions in Brazil until the year 2015.

**Design:** nationwide comparative observational study.

**Patients and Methods:** The population studied encompassed all adult residents ( $\geq 20$  years) living in five Brazilian regions, between 2000 and 2010. Sub-notified deaths were redistributed proportionally to IHD deaths. Age-standardized mortality rates (ASMRs) per 100,000 inhabitants, by sex and region, were calculated, employing a standard Brazilian population, and constructing multivariate regression models to quantify and to project temporal trends.

**Main outcome measures:** Absolute numbers of death due to IHD and the corresponding region-specific death rates in Brazil by age and sex.

**Results :** During the study period, 627,786 men and 452,690 women died due to IHD in Brazil. ASMR trends across all regions for men and women tended to converge, driven by a declining trend in the South and Southeast, and an opposite incline in the North and Northeast ( $p < 0.05$ ). Future projections demonstrated potential widening of the observed north-south gap in coming years.

**Conclusions :** The IHD death trend in Brazil has changed from a decline to a stagnant state. However, a significant discrepancy in mortality trends exists between the northern and southern regions, which is likely to widen further. Re-appraisal of the public health policies tailored to populations with diverse socioeconomic structures, are urgently required.

vital statistics, heart diseases, ischemia, Brazil, disparities.



## INTRODUCTION

Ischemic heart disease (IHD) remains the leading cause of adult deaths in Latin America<sup>1</sup>. Despite overwhelming public health needs, reliable national data on IHD and its temporal trends are generally scarce from this region. Furthermore, Latin America has potentially the world's largest income inequality, which is, as reflected by a *GINI index*, currently higher than for any other region<sup>2</sup>. As economic diversity tends to influence disease severity and therapeutic approaches<sup>3</sup>, it is possible that the scale of the IHD epidemic in Latin America will vary substantially across its regional populations<sup>4</sup>.

Among all Latin American countries, Brazil is considered as having the fastest growing economy and being the most populous nation (comprising almost a third of the entire Latin American population). As other emerging economies such as India and China, Brazil is marked by inequalities in health<sup>5</sup>. Although recent financial reforms have gained the country global recognition and regional leadership, Brazil is yet to overcome major disparities in individual income, education and health care across its regions<sup>6</sup> - factors that may limit considerably the effects of national health care reforms (e.g., the ones aimed at reducing IHD burden)<sup>7</sup>. This is of particular relevance as cardiovascular disease (CVD) is currently the principal cause of death among adults in Brazil<sup>8</sup>. Nonetheless, reliable evidence on the true scale of this burden is conflicting as, despite the rise in CVD risk factors across Brazil<sup>9</sup>, the available literature indicates a gradual decline in IHD-related deaths between 1950 and 2000<sup>10-13</sup>. Interpretation of these earlier reports<sup>10-13</sup> is difficult as they (1) typically lack systematic approaches to reliably quantify the temporal trends of IHD deaths in population subgroups (e.g., age and sex); (2) do not examine findings in the contexts of socioeconomic heterogeneity across regions; and (3) include cause-of-death information collected before the year 2000. The validity of population and mortality data during this period was challenged due to several quality issues (e.g., nonstandardized censuses, significant lack of completeness and inaccuracy of the case definitions). Furthermore, between 1990 and early 2000, Brazil introduced integrated universal health care (*the Unified Health System; SUS*) and cash transfer programs, targeted specifically towards the poor to reduce regional inequality (e.g., *Bolsa Familia Program*)<sup>14</sup>. However, to what extent these reforms may have subsequently influenced the IHD burden, remains unknown. It is also unclear, how the overall Brazilian trend is impacted by potentially discrepant trends across regions.

In this nationwide comparative observational study, we quantified the recent trend of IHD deaths in Brazil during the last decade (2000-2010). We further characterized these trends for various population characteristics (i.e., age group, sex, income, and the regional differences within them). Additionally, based on the records of the past decade, we forecast the upcoming mortality trends across various regions in Brazil up to the year 2015.

## SUBJECTS AND METHODS

### Study design

We conducted a population-wide comparative observational study of mortality from IHD across Brazil. The study encompassed all five government [regions](#) of Brazil as defined by the [Instituto Brasileiro de Geografia e Estatística](#): North, Northeast, Central, Southeast and South regions. The study population involved all adult residents living in these regions, 20 years and older, in each year from 2000 to 2010. The principal outcome measures were absolute numbers of death due to IHD and the corresponding region-specific death rates in Brazil by age, sex and annual trends over time. Since Brazil is ranked medium in the quality of mortality information <sup>15</sup>, it has a proportion of ill-defined underreported deaths that varies across regions. To account for those deaths, we redistributed the proportion of such deaths by age, gender, region and year due to IHD deaths as described elsewhere<sup>16</sup>.

### Data sources

Mortality data were extracted from the Brazilian National Mortality Data System<sup>9</sup> by age group, sex and region of residence. The corresponding information of mean income per capita and per region were available for the years between 1999 and 2009 (except for the year 2000), and were abstracted in collaboration with the National Applied Economics Research Institute in Brazil<sup>17</sup>. Additionally, we included demographic data, for varying age groups and by sex, from the National Census Survey of Brazil<sup>18</sup>. In order to determine the causes of deaths, we used relevant codes for IHD (i.e., I20 to I25) in the International Classification of Diseases or ICD (10<sup>th</sup> Revision), which included angina pectoris, first-ever or acute myocardial infarction (AMI), subsequent [myocardial infarction](#), complications following AMI, and other acute and chronic IHD outcomes.

### *Statistical analysis*

Age-specific mortality rates for every 10-year age stratum and calendar year were calculated, separately for men and women, dividing the total number of deaths by the overall population for each age group in the same year, and subsequently multiplying this number by 100,000 inhabitants, using measures described elsewhere<sup>19</sup>. To allow for comparisons over time and across the regions within Brazil during the study period, age-standardized mortality rates (ASMR) per 100,000 in all age groups were truncated between 20 and 80 years, and were directly standardized employing the Brazilian population in the corresponding year as the standard population. The proportions of early deaths (defined as between 20 and 60 years old) were obtained by dividing the correspondent ASMR by the total (i.e., 20 to 80 years) ASMR in the same year for every region. To compare trends within the regions, a hierarchical multiple linear regression model was constructed, considering region, sex and time as independent variables and ASMR as the dependent variable. For this model, time is nested within region and region is nested within gender. This is presented in the Methodological Supplementary Appendix. The parameters associated with time were estimated for every region, and sex and were used to compare the variation in ASMR between regions for each sex (test of parallelism). Wald's test was used for the comparison. The same estimated model was used for future projections for the years 2011, 2012, 2013, 2014 and 2015, with respective corresponding 95% confidence intervals (CIs), and to estimate the average year-to-year changes (and corresponding 95% CIs) for gender in every region (and overall Brazil). All statistical tests were two-sided, using a significance level of  $p < 0.05$ . All analyses were performed using SPSS 20.0 for Windows (SPSS Inc., Chicago, IL, US) and Stata release 11 (StataCorp, College Station, TX, US).

## RESULTS

### *Distribution and characteristics of the study population*

Demographic and age distributions of the Brazilian population at baseline and in the final year of the study are summarized in Table 1. During the study period, the overall adult population (i.e.,  $\geq 20$  years) increased by approximately 25%. Several individual regions demonstrated small declines in the relative proportion they contributed to the overall Brazilian population from the years 2000 to 2010, such as Northeast (-2.2%), Southeast (-0.6%) and South (-0.4%). By contrast, in the North and Central regions, populations increased by 13% and 3%, respectively. During this 10-year period, the proportion of

those older than 60 years in the overall adult Brazilian population remained almost the same for men and increased marginally for women (2%). Small proportional changes in the population were consistent across all regions. Nonetheless, between 1999 and 2009, the mean income per capita in Brazil increased by 25%. In 2000, the highest income per capita per month in Brazil was recorded in the Southeast region (355.66 US dollars), which was more than double the lowest regional income per capita recorded in the Northeast region in the same year. Although the subsequent within-region increase in the income per capita in the Northeast was higher than in the Southeast (41 vs. 19%), significant economic differences between the northern and southern regions were generally sustained throughout the entire study period.

#### *Burden of deaths from IHD*

Overall, during the entire study period, a total of 627,786 men and 452,690 women died due to IHD in Brazil. In 2000, 52,743 and 38,184 fatal IHD events occurred in men and women, respectively, while 62,854 men and 44,706 women died from IHD in 2010. The latter numbers represented an increase in the absolute deaths by approximately 19% and 17% for men and women, respectively. The highest absolute number of total fatal IHD events in the North, (1,635 vs. 2,895) and the Northeast (9,100 vs. 14,757) for men. Similarly in the Central and Southeast regions, compared to the baseline year, the number of IHD events in 2010 increased by approximately 30% and 10%, respectively. In contrast, the South experienced a reduction (-2.0%) in the number of total IHD deaths. For women, the IHD mortality also increased across all regions, however, varied widely in magnitude. Compared to 2010, fatal IHD as a proportion of all causes of mortality in the whole of Brazil were only marginally reduced in 2010 (<1%).

#### *Age-standardized mortality rates of IHD for different population groups*

For men, the observed ASMRs of IHD in 2010 were 77.9, 93.2, 97.0, 103.7, 108.3 and 94.0 per 100,000 population in the North, Northeast, Central, South, Southeast and overall Brazil, respectively (Figure 1). The corresponding ASMRs for women in the same year were 50.8, 67.6, 61.4, 69.8, 67.7 and 62.8 per 100,000 population. Between 2000 and 2010, the average year-to-year changes in ASMR varied significantly between men and women in the North and Northeast (overall ASMR changes of 2.00 vs. 0.9 and 2.4 vs. 1.8, respectively;  $P < 0.05$  for both) and changes were also marked, albeit negatively, in the South and Southeast regions (overall ASMR changes of -3.0 vs. -2.4 and -1.9 vs. -

1.6, respectively) (Figure 2). In the analyses exploring IHD mortality rates in different age groups between 2000 and 2010, ASMRs were reduced across almost all age groups in the South, Southeast, and Central regions including overall Brazil (data not shown). By contrast, there were generally positive incremental changes across the age groups in the Northeast and North.

#### *Trends in IHD mortality rates across regions between 2000-2010*

In Brazil, the overall IHD mortality trend between 2000 and 2010 remained largely similar with small, borderline significant changes in the ASMRs (-0.84 per 100,000 population, 95% CI: -1.68 to 0,00 for men; 0.48 per 100,000 population, 95% CI: 0.00 to 0.95 for women). In the region-specific analyses, ASMR trends across all regions and genders tended to converge during the study period, driven principally by a declining trend in the South and Southeast, and an opposite incline in the North and Northeast regions (eFigure1). The test of parallelism in these mortality trends yielded significant differences across all regions for men ( $p \leq 0.05$ ). For women on the other hand, the South and Southeast regions showed a non-significant parallel decline ( $p = 0.07$ ), but there were variable significant changes over time for the remaining regions ( $p \leq 0.05$  for all). For the years 2000 and 2010, the magnitude of the ASMRs across all regions was broadly comparable with the levels of their corresponding annual income per capita (eFigure1). Approximately 55% of the overall changes in the ASMR were explained by regional differences ( $p < 0.001$ ).

#### *Forecasts of the overall and regional IHD mortality rates*

Figure 3 illustrates the projections of the IHD mortality rates up to the year 2015 based on the current observed trends, separately by region and sex. For men, baseline (2000) and the projected end year (2015) ASMR of IHD per 100,000 residents were 134.4 and 88.3 for the South and 127.6 and 98.7 for Southeast regions. By contrast, the ASMRs were 58.1 and 87.9 for the North and 69.0 and 105.4 for the Northeast, respectively. These within region ASMRs were similar in women. The corresponding ASMRs for all of Brazil remained broadly unchanged between 2000 and 2015. For both men and women, during the study period, there was an apparent reversal of mortality trends across several regions.

## DISCUSSION

Overall, the age-standardized mortality rates of IHD in Brazil remained broadly consistent between 2000 and 2010 for all age groups and genders and over time. Nonetheless, these observed rates differed considerably across individual regions of disparate economic structure, with distinct opposite mortality trends between the northern and southern parts. Additionally, if these discrepant region-specific trends of IHD death rates were to continue, there would be a further widening of this existing discrepancy within Brazil by the end of 2015, according to future projections.

Our findings may have several explanations. Although a rapid increase in purchasing power has been a general feature across all Brazilian regions in the recent past, IHD mortality rates seem to have increased in the comparatively poorer North and Northeast, and are in decline in the wealthier South and Southeast regions. It is possible that these regions might have undergone distinct epidemiological phases at the same time. As CVD risk factors tend to be distributed more disproportionately among the poorer population subgroups, this may potentially explain the increasing pattern of IHD deaths in the northern region over time. Furthermore, deeper differences in demographic and lifestyle factors may also determine excess mortality in the northern part. For instance, it has been previously reported that, compared to the southern counterparts, northern regions in Brazil have significantly lower human development indices and lower literacy rates. Similar disparity in the regional health care structures and public health resources may further reinforce such differences. There are fewer sophisticated cardiology facilities, limited options or affordability in accessing private health care, and crucially, less per capita expenditure on health care available in the northern regions <sup>20</sup>. Despite recent improvements in overall life expectancy (owing largely to changes in social health determinants<sup>21</sup>), existing diversity in socioeconomic status, if sustained, will continue to challenge the goal to achieve equitable universal health care in Brazil<sup>20</sup>. Interestingly, IHD mortality gradient across regions is also described in other developing nations<sup>22 23</sup>.

Nationally, no further reduction in overall IHD mortality was evident thus far for either men or women. Potentially different and overlapping stages of epidemiologic transition across Brazil may have led to a counteractive cycle in the populations ageing process and poverty induction<sup>24</sup>, which coupled with the ever widening regional health disparities may have undermined the efforts to strengthen the national health care system. This is of particular concern, as Brazil appears to have one of the worst national ASMR statistics for

IHD in South America, only surpassed by Colombia and Venezuela<sup>25</sup>. Finally, it is also possible that over time many healthier people from the poorer regions may have migrated to wealthier health regions due to better opportunities. If true, such “selective migration”<sup>26</sup> may further explain the higher IHD death rates in the northern parts.

The findings and methods in this study differ from those of previous reports in several ways. Previous studies have consistently described overall declines in the IHD death rates in Brazil during the last decades across all ages and for both men and women<sup>10, 27-28</sup>. These investigations, however, involved only partial assessments of mortality based principally on the capital cities or on smaller geographical areas. By contrast, our results present a more recent and comprehensive quantification of the population risk at the region level. As opposed to the earlier studies, which typically present more qualitative measures (e.g., standardized mortality ratios or univariate death rates), we used multivariate linear regression models for specific quantitative comparisons involving age group, sex, time and region structures as independent covariates.

Our findings may have important implications. First, they extend the previous literature by demonstrating a more static, yet considerably high overall IHD mortality rate in Brazil. Second, they indicate that the health gap may have widened between the northern and southern regions, increasing steeply from 2000 to 2010 despite new economic and health reforms to reduce inequalities in Brazil. They highlight that all future public policy emphasis targeting economically diverse Brazil (and other South American populations) should be region-specific. Third, they support the recently launched bespoke national plan for effective monitoring, health promotion and integrated care of the NCDs. However, as the overall trend is likely to be heavily influenced by regional disparities, the current target outlined (i.e., annual reduction of NCD-related deaths by 2%) would require re-appraisal to reflect the current differences in sociodemographic factors and vascular burden. Fourth, these results appear to be in line with the most pessimistic of the three future global scenarios (baseline, optimistic and pessimistic) projected by World Health Organization to provide a wealth of information about possible patterns of global deaths between 2002 and 2030<sup>29</sup>. Therefore, they stimulate further scientific and innovative health policy research (e.g., reduction of vascular risk factors by appropriate local interventions), tailored to the needs of a heterogeneous population.

The strengths and limitations of this study merit considerations. To the best of our knowledge, this is the first study in Brazil to assess the overall trend of IHD in the context

of regional differences. We employed robust analytical approaches to quantify mortality for different population groups living in diverse economic circumstances. We used the Brazilian population as the standard to describe trends in different regions at the same time, which may reduce distortion of the overall findings<sup>30</sup>. Nonetheless, these findings may have been limited, at least in part, by the use of secondary data. Although the overall survey quality in Brazil has significantly increased since 2000 and our correction of underreported deaths as IHD deaths, a residual harmonization of the overall data collection process could vary across areas<sup>31</sup>, which may potentially increase the regional discrepancies. Furthermore, as the mortality database in Brazil is yet to be linked with the primary, secondary or tertiary clinical care information, it was not possible for us to assess any potential interplay by specific treatment measures on the observed mortality trend. We were unable to address residual confounding by any “unmeasured” factors that could additionally affect the precision of our estimates. Finally, to what extent these findings could be applied to other related cardiometabolic conditions (e.g., stroke, heart failure and diabetes) is unclear, and this uncertainty would, therefore, require further characterization.

In conclusion, available data indicate that the IHD mortality trend in Brazil has changed from declining to a stagnant state in the last decade. However, a significant gap in the mortality trend exists between the northern and southern regions, and given the current circumstances, it could widen in the near future. Given that the prevailing socioeconomic diversities may further influence overall mortality and level of health attained, future national health policies and programs in Brazil should reflect these overwhelming disparities and trends to prevent cardiovascular disease in the coming years.

Contributorship statement- Cristina P. Baena and Rajiv Chowdhury participated in the study conception, acquisition, analysis and interpretation of data, writing and reviewing the manuscript. José Rocha Faria-Neto and Oscar H. Franco participated in the conception, appraisal and interpretation of data, and critically reviewed the manuscript drafts. Nicolle Schio and Luiz Cesar Guarita-Souza contributed to data gathering, analysis, interpretation, and critically reviewed the manuscript drafts. Marcia Olandosky and Ary Elias Sabbag contributed to design, analysis, interpretation and critically reviewed the manuscripts drafts.



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Table 1: Distribution of Brazilian population by sex, age and region in the total adult population and mean per capita income

Year	% total >20 y population**	Monthly mean per capita income\$	Men				Women				
			20-60 y		>60 y		20-60 y		>60 y		
			n*	(%)	n*	(%)	n*	(%)	n*	(%)	
<b>Brazil</b>											
2000	100	281.7 ‡	425	(86.6)	65	(13.3)	445	(84.7)	80	(15.2)	
2010	100	383.5	523	(85.1)	92	(14.9)	549	(82.7)	114	(17.2)	
North											
2000	4.5	204.7 ‡	30	(89.3)	4	(10.7)	29	(89.1)	4	(10.8)	
2010	5.0	247.0	41	(88.4)	5	(11.5)	41	(88.2)	5	(11.8)	
Northeast											
2000	20.7	157.0 ‡	108	(85.5)	18	(14.4)	116	(84.1)	22	(15.1)	
2010	18.5	229.3	137	(84.8)	24	(15.1)	147	(82.9)	30	(17.0)	
Central											
2000	4.7	301.6 ‡	30	(88.8)	4	(11.1)	31	(88.8)	4	(11.1)	
2010	5.1	467.8	40	(87.0)	6	(12.9)	41	(86.4)	6	(13.5)	
Southeast											
2000	31.4	355.6 ‡	191	(86.6)	29	(13.3)	201	(84.1)	38	(15.8)	
2010	30.8	471.6	228	(84.6)	41	(15.3)	240	(81.6)	54	(18.3)	
South											
2000	10.7	322.2 ‡	66	(86.5)	10	(13.5)	68	(84.2)	13	(15.7)	
2010	10.4	459.9	77	(84.1)	15	(15.8)	80	(81.4)	18	(18.5)	

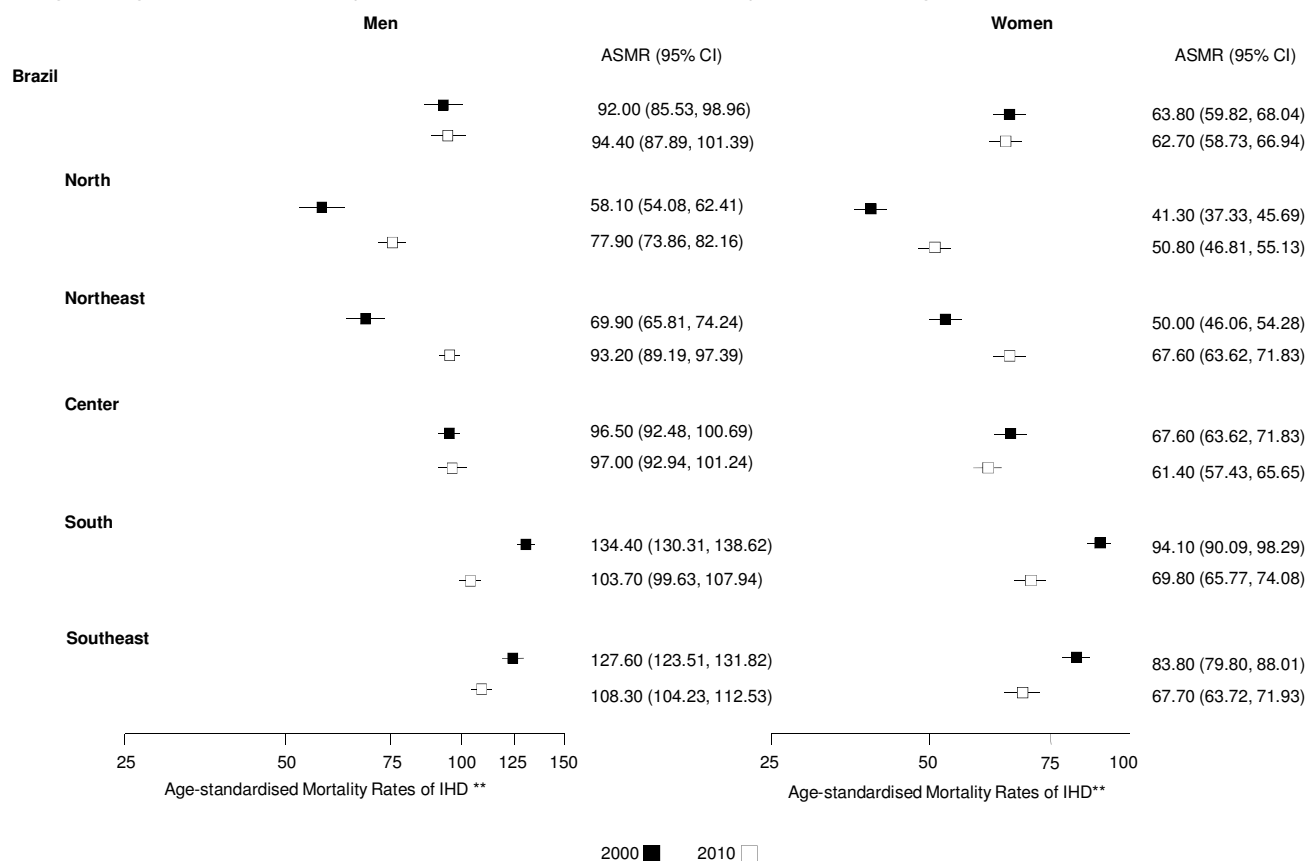
\* population frequencies divided by 100,000

\*\* Proportion of the regional adult population relative to the country's adult population.

\$ Monthly mean per capita income \$ in US dollars. Brazilian currency – 1 Real equivalent to 0.5 in international dollars (09/09/2012)

‡ data from 1999

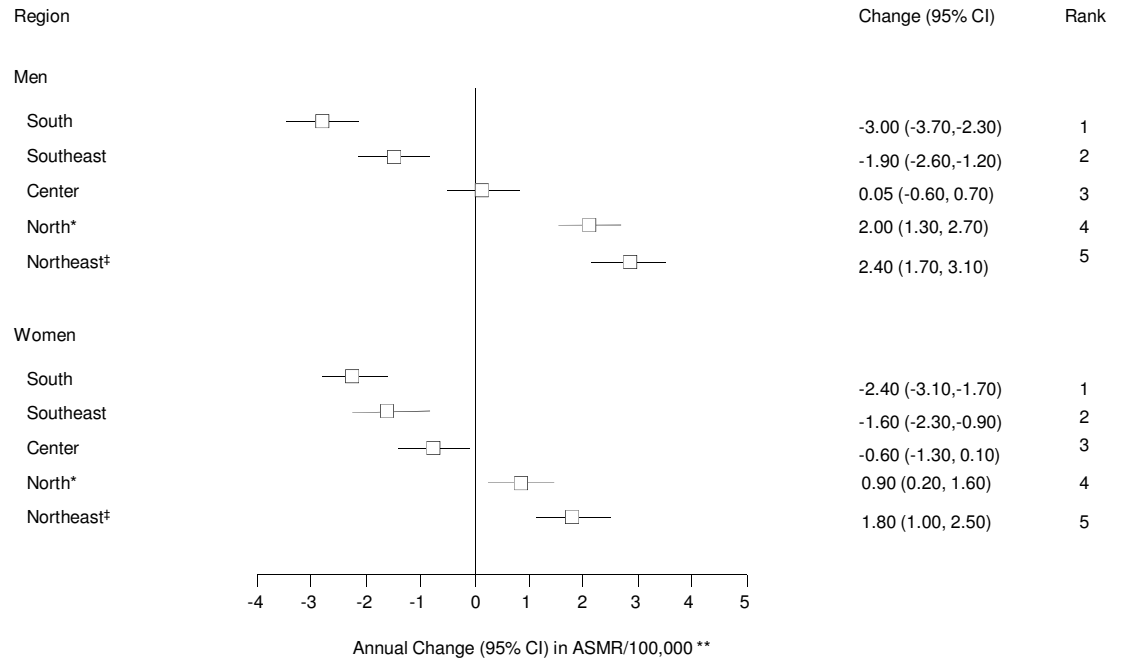
Figure 1: Age-standardized Mortality Rates (ASMR) of Ischemic Heart Disease by sex in different regions across Brazil \*



\* hierarchical multiple linear regression model.

\*\* age-standardised mortality rates were plotted on a logarithmic scale

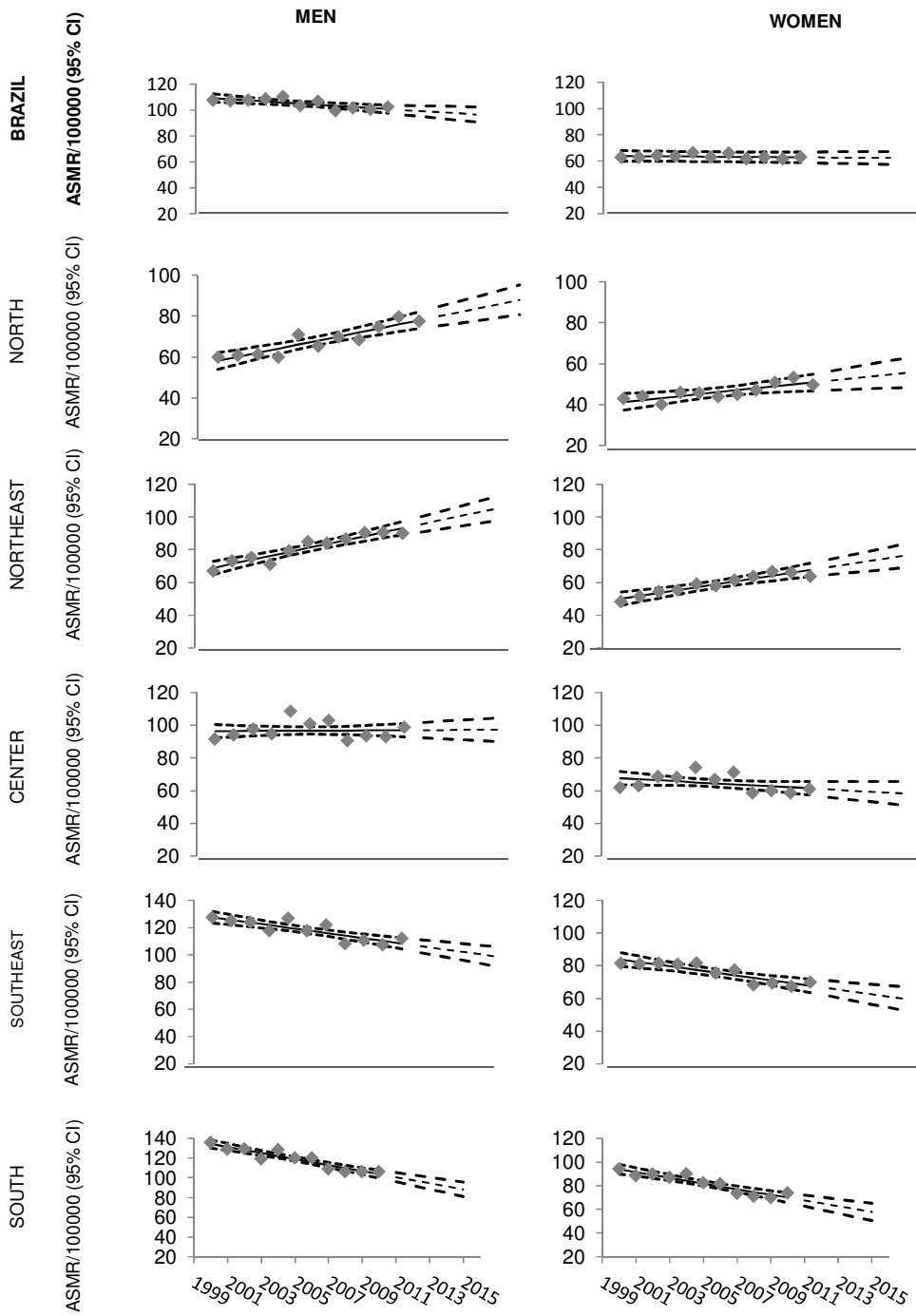
**Figure 2: Changes in the Ischemic Heart Disease mortality by sex in different regions across Brazil (2000-2010)**



\*P <0.05 for difference between men and women; †P <0.05 for difference between men and women

\*\* hierarchical multiple linear regression model

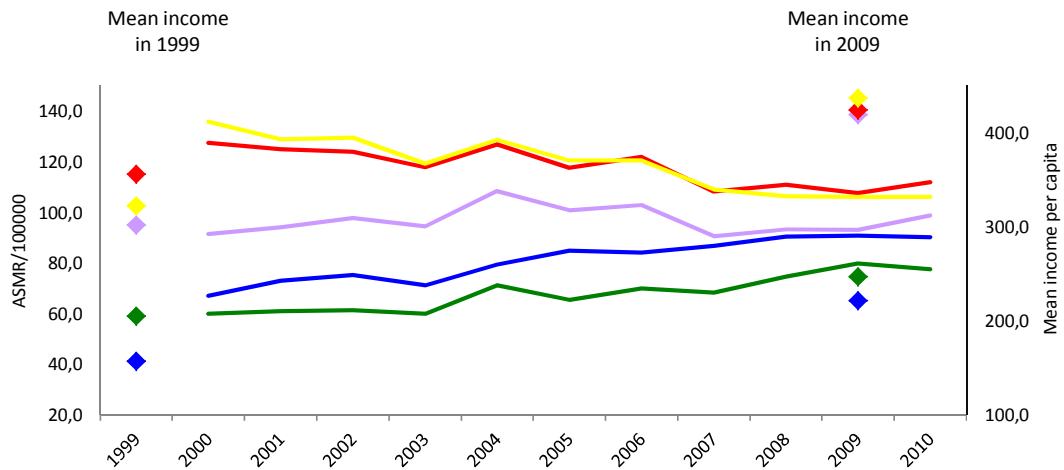
Figure 3: Projections for the IHD mortality rates in Brazil up to 2015 by region and sex



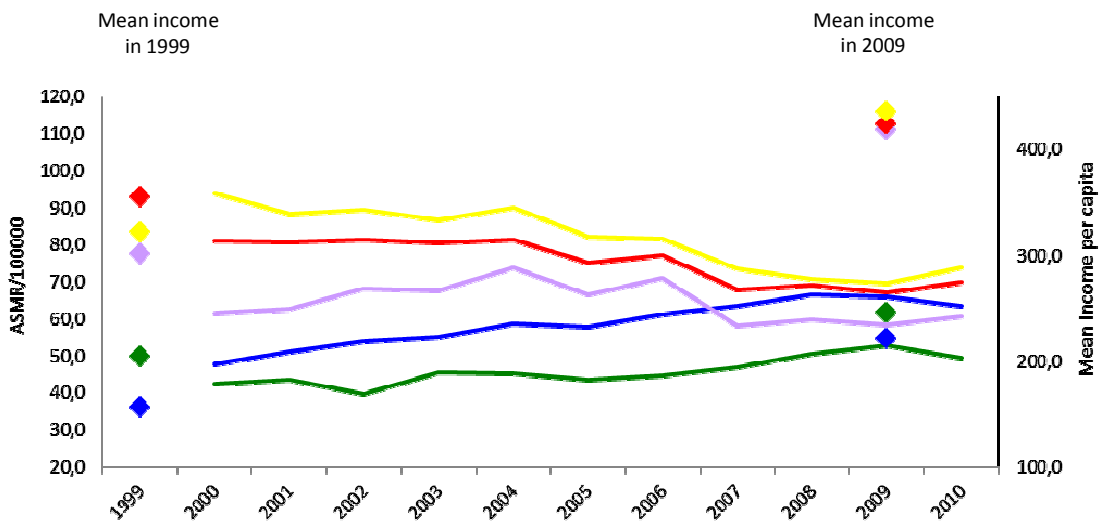
\* hierarchical multiple linear regression model

**efigure :** Trends in IHD mortality rates in Brazil by sex and regional per capita average income

**(1) Men**



**(2) Women**



ASMR                    — N    — NE    — C    — SE    — S  
 Mean income        — \$ N   — \$ NE   — \$ C   — \$ SE   — \$ S



### Methodological Supplementary Appendix:

To compare trends with in the regions, a multivariate linear regression model was constructed where ASMR was entered as the dependent variable, and sex, region and time were entered as independent variables. Appropriate “dummy” variables were constructed for the regions. The final model which was employed to compare the trends is expressed as the following:

$$y_{ijt} = \mu + \beta z_i + \sum_{j=1}^4 \gamma_{ij} + \sum_{j=1}^5 \delta_{ij}t + \varepsilon_{ijt}$$

Where,  $y_{ijt}$  represents the ASMR for sex ( $i=1$  for male and  $i=0$  for female), in the region  $j$ , in year  $t$ , respectively;

$z_i = 1$  if  $i=1$  (male) and 0 if  $i=0$  (female)

$\gamma_{i1} = 1$  if region North, and 0 if not;

$\gamma_{i2} = 1$  if region Northeast, and 0 if not;

$\gamma_{i3} = 1$  if region Central, and 0 if not;

$\gamma_{i4} = 1$  if region Southeast, and 0 if not;

(Note: South is attributed 0 in each of the 4 dummy variables from other regions)

$\delta_{ij}$  represents the ASMR change by every year, for sex  $i$  in region  $j$ ; and

$\varepsilon$  represents the Error. |

## **ARTIGO 3**

### **Effects of lifestyle-related interventions on blood pressure in low and middle income countries: systematic review and meta-analysis**

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**Submetido**

## **Effects of lifestyle-related interventions on blood pressure in low and middle income countries: systematic review and meta-analysis**

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**ABSTRACT** (word count: 291)

**Background** Despite a plethora of evidence supporting the efficacy of antihypertensive agents, hypertension remains poorly controlled worldwide, and in low and middle income countries (LMICs) in particular. In this context, whether non-pharmacological lifestyle-based interventions are effective to help maintain optimum blood pressure levels in the LMIC, is unclear.

**Methods and Findings:** We performed a systematic review and meta-analysis of available clinical trials, based in the LMIC and published before June 2012, which assessed effects of any lifestyle-related intervention on blood pressure. We identified relevant studies by systematically searching multiple electronic databases. Interventions were grouped into behavioral counseling, dietary modification, physical activity and multiple interventions. We calculated and pooled effect estimates (weighted mean difference or WMD in mm Hg between the post- and pre-intervention measurements) with random-effects models. Where appropriate, subgroup and meta-regression analyses were employed to evaluate origins of heterogeneity. From the initial 6211 references identified, 52 studies were included, comprising a total of 6,779 non-overlapping participants. The changes (and corresponding 95% CI) in the WMD expressed as mm Hg achieved for systolic blood pressure were -5.39 (-10.73, -0.05) for behavioral counseling, -3.48 (-5.45, -1.50) for dietary modification, -11.37 (-16.06, -6.68) for physical activity and -6.09 (-8.87, -3.32) for multiple lifestyle interventions. Findings were generally consistent for diastolic blood pressure. Nonetheless, heterogeneity was generally high across these studies, which was partly explained when studies were grouped according to sample size, duration of follow-up and analytical approach used.

**Conclusion:** Available data indicate that lifestyle-related interventions are effective in lowering blood pressure in LMIC. The potential of cost-effective, easily scalable lifestyle interventions is attractive in resource-poor settings as a complementary approach to help shape preventive guidelines. Nonetheless, further investigations with sufficient power and scientific rigor would be required to more reliably quantify these effects.

## INTRODUCTION

Cardiovascular disease (CVD) remains the leading cause of death and disability worldwide particularly in low and middle income countries (LMIC)<sup>1 2</sup>. Although improvements in prevention and care among high income countries have resulted in large reductions of CVD mortality in the last decade, in LMICs, the prevalence of the established CVD risk factors at all ages and the CVD disease burden are still on the rise<sup>3</sup>. Blood pressure remains one of the major modifiable vascular risk factors<sup>4</sup>. Aetiology, prevention and management of high blood pressure have been widely studied<sup>5, 6</sup>. However, high blood pressure remains poorly controlled and the general awareness among individuals and populations with hypertension, especially in the LMICs, is considerably low<sup>7, 8</sup>. This is of particular importance as high levels of blood pressure continues to be a top attributable cause of overall death worldwide, contributing approximately 14% of all global premature deaths<sup>1 9</sup>. Furthermore, 80% of the global burden of death attributable to high blood pressure occur in LMIC<sup>10</sup>- highlighting the urgent need for more intensive efforts to reduce the disease and mortality burden secondary to hypertension in these settings<sup>11</sup>.

To alleviate such overwhelming burden associated with blood pressure, effectiveness of drug therapy has been tested and demonstrated in the LMI countries. However, its availability, affordability and, therefore, adherence to treatment poses a significant challenge<sup>12 13</sup>. By contrast, safe, cost-effective and easily scalable lifestyle interventions are generally considered the first choice to prevent and control optimum blood pressure. Despite huge potentials offered by these interventions particularly for the LMI nations, evidence to support a benefit to effectively reduce blood pressure in the LMIC is scarce and inconsistent. Currently the majority of guidelines for CVD prevention in LMICs rely on evidence gathered mainly from populations derived from high income countries. How applicable they are for the local specific populations remains unknown<sup>14 15, 16</sup>.

We report a systematic review and meta-analysis to summarize and quantify the available evidence on the effects of lifestyle-related interventions conducted in the LMI countries. Additionally, we have evaluated whether the effects may vary across a wide range of study-level characteristics including ethnicity, types of interventions, geographical location and age.

## **METHODS**

Our systematic review was conducted with a predefined protocol in accordance with PRISMA guidelines and an extension of the CONSORT statement<sup>17</sup> to trials evaluating non pharmacological treatments<sup>18,19</sup>. We conducted electronic searches through Medline-Pubmed, Embase, Cochrane Library, CINAHL, Web of Science, Scopus, Scielo and LILACS searching for suitable references published between January 1977 and September, 2012 (date last time the search was run). We used combinations of Medical Subject Headings (MeSH) and free text words that included search terms related to the population (eg *developing country or low and middle income country*) and lifestyle interventions, which were combined with search terms related to the outcomes (eg blood pressure, hypertension). The detailed search strategy and countries included are presented in **Appendix 1** and **2**. We sought for Randomized Clinical Trials, Clinical trials and Community trials. If data were unavailable, or when local libraries could not retrieve the full paper, authors were contacted by e-mail. No language restriction was applied and studies were translated by native speakers experienced in the health field. In the case of multiple publications, the most recent and complete report was included.

### *Study selection*

Two authors independently reviewed each title and abstract to determine whether the study met the inclusion criteria. We sought for studies with a parallel design evaluating non-pharmacological (eg dietary modifications, physical activity or educational/behavioral counseling) interventions among adult populations from LMICs, reporting blood pressure levels (systolic and/or diastolic) in mm Hg at baseline and after intervention in the same sample. We excluded studies if they

were: conducted solely in developed (high income) countries according to the World Bank incoming group list used by WHO (**Appendix 1**); cost effectiveness studies; assessed effects of a pharmacological intervention; involved pregnant women; animal studies and, if the intervention lasted less than 8 weeks<sup>20 21</sup>. Any disagreements in article selection were solved through discussion and if necessary a third author was consulted to resolve disagreement. Full text articles were retrieved for the selected titles after initial appraisal. In order to find unidentified studies with the search run in databases, we checked reference lists, contacted authors directly and conducted citation tracking. Articles selected were retrieved and assessed by two independent authors to ensure they met full inclusion criteria. Any disagreements were resolved through discussion with a third independent reviewer.

#### *Data extraction*

Two independent reviewers extracted data using a pre-designed data collection form on study and participant characteristics, comparison groups, outcomes, analysis and conclusions. Study characteristics recorded included: author, year of publication, geographic origin and setting of the study, design and funding source, gender of participants, ethnicity, age, duration of intervention, residence, drug treatment, co morbidities, inclusion and exclusion criteria, type and delivery of intervention, quality, intention to treat analysis, tests used to analysis, participants in the control and intervention group, blood pressure measurements before and after intervention in both groups with respective variances. We translated physical activity interventions, into metabolic equivalents (METs) as a co-variate of interest<sup>22</sup>. Multiple interventions were defined as more than one of the above interventions delivered at the same time compared to regular care or no intervention. We recorded the pre and post intervention blood pressure measurements in mm Hg and the variance as standard deviation, standard errors or 95% confidence interval as reported. In the case of multiple publications from the same study, the most comprehensive or recent information, as appropriate, was included in the analysis. In order to avoid repeating the same control group in the analysis, if the same

study presented more than one intervention arm comparing to the same control group, we only included the intervention with the largest sample in the case of diet interventions, and the highest intensity arm group in the case of physical activity. Two independent reviewers evaluated possible bias and intention-to-treat in individual studies according to the Cochrane Tool for Bias Assessment<sup>23</sup>.

### *Statistical Analysis*

Both systolic and diastolic blood pressure measurements were noted as a continuous variable in mm Hg. The effect sizes were summarized as weighted mean difference between post- and pre-intervention measurements in each trial<sup>24</sup>. If data were unavailable, we used standard equations (**Appendix 3**) to impute the missing variance. We employed Der Simonian and Laird random effect models with a 95% CI<sup>25</sup> to perform the meta-analysis. A random-effects model was chosen to conservatively deal with potential heterogeneity among the studies included. Heterogeneity among studies was measured with I-squared<sup>24</sup>. We divided studies according to type of interventions (behavioral counseling, dietary modification, physical activity and multiple interventions). To evaluate potential differences between specific subgroups we repeated the analyses by key factors. We also conducted metaregression (maximum likelihood method) to identify the key sources of heterogeneity<sup>26</sup>. Publication bias was examined by using funnel plots (plotting sample size against effect size) and non-parametric trim-and-fill methods<sup>27</sup>. Any potential small studies effects were investigated by Egger's test<sup>27</sup>. Additionally, to investigate the influence of a single study on the summary effect, we performed sensitivity analyses in the four groups of different interventions<sup>24</sup>. All analyses were conducted using Stata version 12.0 (StataCorp LP, College Station, Texas, United States).



## RESULTS

### *Identification of the relevant studies*

From the initial 6211 references identified, 409 were considered potentially eligible and were retrieved. The reasons for exclusion of references to quality assessment are shown in **Figure 1**. Of the 52 studies included in the systematic review, 43 were eligible for the meta-analysis including 6,779 participants.

The reasons for exclusion of nine papers were control group and intervention group receiving similar interventions<sup>28 29</sup>, interventions were at organizational level<sup>30</sup>, intervention involved pharmacologic supplements or drug<sup>31 32</sup> with lifestyle, the results could not be pooled due to lack of variance in pre intervention and post intervention<sup>33, 34 35</sup> or doubled population<sup>36</sup>.

Since studies differed in terms of type of interventions, they were divided in 4 main groups according to delivery (eg. physical activity, behavioural counselling, dietary modifications, and multiple interventions). The random effect sizes and heterogeneity are shown in **Figure 2**.

### *General characteristics of the included studies*

The characteristics of the populations and interventions included in the systematic review are shown in **Table1**. The studies included in the meta analysis involved 6.779 participants (44,5% women; mean age 53.5±9.5 years, range 18-80 years). The duration of the studies ranged from 2 to 30 months, with a mean duration of 5.3±4.8 months. Approximately half of the studies (23 studies) reported co-morbidities and the most common were,diabetes<sup>35, 37-45</sup> and metabolic syndrome<sup>46-48</sup>.

### *Effects of various lifestyle interventions on blood pressure*

The differences were significant overall. When all studies were pooled irrespective of the type of intervention, the overall effect CI (95%) on systolic blood pressure was -6.46 (-8.31, -4.60) mm Hg and for diastolic blood pressure -3.68 (-4.74, -2.62)mm Hg. This effect and the heterogeneity differed by type of interventions

**(Figure 2).** Thirty-three studies were randomized controlled trials<sup>38, 42, 44-74</sup>, two were community trials<sup>41, 75</sup> and eight were clinical trials<sup>37, 39, 40, 43, 45, 76-78</sup>.

We included 14 studies reporting physical activities interventions, involving 1014 participants (52.6% women; mean age 59.2±7.5 years). The duration varied from 2 to 12 months with mean of 3.7±2.3 months. The studies were conducted in Brazil<sup>37, 39, 43, 59, 62, 76, 78</sup>, China<sup>56, 65, 67</sup>, India<sup>44</sup>, Thailand<sup>69</sup> and Nigeria<sup>72</sup>. When pooled, physical activity interventions lowered blood pressure by -11.37 (-16.06, -6.68) mm Hg for systolic pressure and -6.54 (-9.49, -3.59) mm Hg for diastolic blood pressure. Subgroup analysis yielded larger results from studies with sample size smaller than 100 participants (**Figure 3**). The differences in effect sizes for smaller samples were -4.82 (-4.91, -4.72) mm Hg for systolic blood pressure and -5.58 (-5.6, -5.5) mm Hg for diastolic pressure. Other key subgroups effect sizes are shown in **Table 2**. Metaregression did not indicate an influence of gender, (P >0.05 for both systolic and diastolic blood pressure). Mean age showed a borderline significant effect (P=0.057) on systolic blood pressure and no influence on diastolic blood pressure (P >0.05).

We included 8 studies on behavioral counseling in the meta-analysis with the total population of 1,831 participants (40.3% women; mean age 53.8±6.6 years). The interventions were conducted in China<sup>55, 73, 75</sup> Turkey<sup>42, 49</sup> India<sup>64</sup> Jordan<sup>38</sup> and South Africa<sup>61</sup>. Hypertension was reported in 4 studies<sup>38, 40, 42, 64</sup>. Studies varied in duration from 2 to 30 months with a median of 7 months. Information available in the studies allowed for subgroup analysis, shown in Table 3. The pooled results for behavioral counseling showed a reduction of -5.39 (-10.73, -0.05) for systolic pressure and -2.55 (-4.99, -0.11) for diastolic pressure. Effect sizes in samples smaller than 100 participants were -5.39 (-10.73, -0.05) mm Hg and -2.55 (-4.99, -0.11) mm Hg lower than bigger samples for systolic and diastolic pressure respectively (**Figure 3**). Key factors subgroup effect sizes are shown in **Table 3**. Mean age influenced the effect size across trials (P=0.03) by 0.8 mm Hg 95%CI (1.2, 0.35) decrease in systolic blood pressure to one year increased in mean age but it did not influence diastolic blood pressure effects sizes (P=0.26). Proportion

of women did not influence systolic or diastolic blood pressure effect sizes ( $p=0.78$  and  $p=0.86$ , respectively).

We included 13 studies reporting dietary modification in which, 1,831 participants were involved (53.4% women; mean age  $47.3\pm 8.8$  years,). The studies were conducted in China<sup>45-47, 51, 70</sup>, Brazil<sup>57, 68, 71, 74</sup>, South Africa<sup>48, 63, 66</sup>, India<sup>55</sup> and Iran<sup>48</sup>, Three studies enrolled only women<sup>46-48</sup>. Mean duration was  $4.3\pm 2.7$  months. Comorbidities reported were metabolic syndrome, diabetes<sup>45, 51</sup> and hypertension<sup>45, 51, 70</sup>. The pooled results for dietary interventions were  $-3.48$  ( $-5.45$ ,  $-1.50$ ) mm Hg for systolic pressure and  $-2.25$  ( $-3.61$ ,  $-0.89$ ) mm Hg for diastolic pressure. Studies involving smaller samples showed larger results and the difference was  $-1.56$  ( $-1.60$ ,  $-1.51$ ) mm Hg for systolic pressure and  $-1.92$  ( $-1.95$ ,  $-1.88$ ) mm Hg for diastolic pressure (**Figure 3**) in studies involving less than 100 participants. We analyzed different types of diets and their effects on systolic and diastolic blood pressure: mineral replacement diets (3 studies, I-squared= 0%, 59%), grains and fibers rich diet (1 study), protein rich diets (4 studies, I-squared= 73.2%, 83.6%), nuts and seeds rich diets (2 studies I squared = 0%, 0%) and complex patterns diets (3 studies, I-squared= 0%, 0%). The effects sizes are shown in **Table 4**.

We included eight multiple interventions studies enrolling 2103 participants (46.3% women, mean age  $53 \pm 11$  years). These studies were conducted in India<sup>58</sup>, China<sup>50</sup>, Costa Rica<sup>41</sup>, Iran<sup>52</sup>, Turkey<sup>53</sup>, Chile<sup>54</sup>, Malaysia<sup>60, 77</sup> and Argentina<sup>77</sup>. Coronary heart disease was reported in 3 studies<sup>35, 41, 53</sup>. Mean duration of the studies were  $5.5 \pm 3.2$  months. The studies combined physical activity and diet or behavioral counseling interventions. Pooled effect sizes were  $-6.09$  ( $-8.87$ ,  $-3.32$ ) mm Hg for systolic blood pressure and  $-2.43$  ( $-3.74$ ,  $-1.12$ ) for diastolic blood pressure. The combination physical activity and diet or physical activity and/behavioral counseling did not yield significant difference in effect sizes ( $p > 0.05$ ). In this group, results in sample sizes smaller than 100 participants differed by  $4.67$  ( $4.62$ ,  $4.71$ ) mm Hg to systolic pressure and  $-3.74$  ( $-3.76$ ,  $-3.71$ ) mm Hg

from the results in studies with bigger samples (**Figure 3**). Information about the intensity of physical activity did not allow for transformation to METs. These findings were further investigated in various study-level subgroups (**Table 5**). Meta-regression analyses indicated that the proportion of women did not influence effect sizes for either systolic or diastolic blood pressure (Metaregression  $p > 0.05$  for both measures). Mean age also did not influence effects in systolic or diastolic blood pressure ( $p > 0.05$ ).

## DISCUSSION

Overall, we found that lifestyle interventions of any type (behavioral, physical activity or dietary interventions, given singly and in combination) significantly lowered blood pressure levels in the LMI settings by approximately -6.46 mmHg and -3.68 mmHg for systolic and diastolic blood pressure, respectively. Effects ranged from -11.37 mmHg for physical activity to -3.48 mmHg in dietary modification for systolic blood pressure and from -6.54 mmHg for physical activity to -2.25 mmHg for dietary modifications for diastolic blood pressure. Nonetheless, heterogeneity was generally high across the available studies, which was partly explained when studies were grouped according to sample size, duration of follow-up and analytical approach used.

In the current review, when assessed according to intervention type, interventions that promoted physical activity lowered systolic and diastolic blood pressure significantly, irrespective of antihypertensive drugs usage, ethnicity, mean age and randomization of patients. Interestingly, interventions lasting 2 to 4 months showed a significant lowering effect on systolic and diastolic blood pressure than the ones lasting 4 to 12 months, reinforcing results in a previous meta-analysis on physical activity<sup>79</sup> where longer lasting effects were highlighted as a challenge. Physical activity interventions involving more than 100 participants did not show a significant lowering effect on diastolic blood pressure. Similarly, resistance exercise programs did not show a significant lowering effect to systolic blood pressure and this subgroup findings differ from a previous meta-analysis on resistance program exercise effect to blood pressure<sup>80</sup>. Different inclusion criteria and our reduced number of studies in this subgroup may explain the difference in results. Additionally, interventions offering less than 2 hours per week of physical activity did not show a significant effect on systolic or diastolic blood pressure. The large difference found between those effect sizes could be due to lower quality of those studies reporting less physical activity. These findings are in line with the first meta-analysis on the lifestyle interventions effect to blood pressure we could find<sup>81</sup>

suggesting the difficulties of sustaining exercise routine. Comparison of our study with other meta-analysis is impaired by the lack of other reviews focusing on low resource sensitive settings.

For both physical activity and behavioral counseling interventions, the effect sizes were larger in the studies that reported substantial co-morbidities –indicating that high risk individuals may have larger benefits from the lifestyle interventions. Conversely, effects were not significant in studies reporting blinding of outcome assessors and intention-to-treat analysis which potentially points out quality issues in this type of intervention. Behavioral counseling interventions were effective in lowering systolic and diastolic blood pressure in studies reporting anti-hypertensive drugs and in studies with follow up from 2 to 6 months. These findings might have been driven by the nature of most of the behavioral counseling studies, which involves adherence to medication program. Furthermore, they might also indicate the challenge of sustaining a healthy lifestyle since studies with longer duration showed a significantly lower effect size as was described for physical activity interventions as well.

Studies on dietary modifications, labeled as complex dietary pattern comprised Dietary Approach to Stop Hypertension and Mediterranean diet (DASH)<sup>82</sup>, and showed a significant lowering effect on systolic and diastolic blood pressure. Mineral replacement based diets (e.g. potassium replacement) studies showed significant effect in lowering systolic, but not diastolic, blood pressure. These two groups showed no heterogeneity even though the interventions were conducted in 3 different continents. These findings might indicate a consistent lowering effect of mineral replacements diets on systolic blood pressure. Nuts and seeds based diets and protein rich diet studies did not yield significant effects on systolic, nor on diastolic pressure. Subgroup analysis was hampered within the dietary modifications due to reduced number of studies and the low power we found after grouping studies according to different dietary patterns.

Multiple interventions group was composed of interventions involving physical activity combined with nutrition training or behavioral counseling. There was a consistent lowering effect on systolic and diastolic blood pressure regardless of ethnicity, duration of the intervention, reporting of co-morbidities and intention-to-treat analysis. In studies reporting blinding of the outcome assessor, the reduction was significant for systolic and diastolic blood pressure, but not in studies in which there was unclear reporting of the blinding. Effects on Asian participants were greater than in South Americans and studies lasting 2 to 6 months showed larger effect sizes than longer studies, especially for diastolic blood pressure. Multiple intervention studies involving less than 100 participants showed no heterogeneity across them suggesting a consistent effect from those interventions.

We observed considerable heterogeneity among the included studies which were partly explained by our subsidiary subgroup analyses involving differences in sample sizes, duration of intervention, blinding of outcome assessors and intention to treat analysis. On the other hand, multiple interventions studies seemed to show the opposite effect on smaller samples, with very low heterogeneity among the smaller sample study group. This indicates a consistency in the interventions involving physical activity and nutrition training or behavioral counseling at the same time involving samples smaller than 100 participants in LMIC settings. Possible explanations for these findings could be the close supervision given to smaller samples and higher quality of reports in those interventions.

Our results somewhat differ from previous meta-analyses of the effects of lifestyle interventions on blood pressure<sup>21 83</sup>. We found larger effect sizes for physical activity interventions and for multiple interventions than others did. Nevertheless, we found similar larger effects in participants using antihypertensive medications in the cases of physical activity and behavioral counseling<sup>83</sup>. Our results on dietary modifications were in smaller magnitude than previously reported although other meta-analysis showed the same direction of effect. Some differences in effect sizes might be due to different selection criteria on study designs as some other

authors included cross over studies whereas we included only parallel studies. Including pharmacological supplementation (e.g. potassium, magnesium, calcium and fish-oil) and separated sodium restriction diet from diet interventions while we included sodium restriction on dietary interventions might also explain different results. Additionally, our comparison with other reviews is impaired by the difference in population of interest as we focused on LMIC populations. This might be reflected by the fact that we found very few references included in our study found in other meta-analysis mentioned in this discussion.

Strengths and limitations of the current study merit consideration. We were unable to evaluate the associations in the contexts of potentially important demographic information such as living in rural or urban areas and educational status as most studies did not provide these data. Despite the considerable number of references included in this meta-analysis, an apparent lack of large-scale, high-quality trials in the LMI countries. Nonetheless, to the best of our knowledge, this is the first quantitative synthesis of all available intervention studies based on LMIC and assessing effects of non-pharmacological lifestyle interventions on blood pressure. The fact that the majority of our included studies were not included in the previous meta-analyses suggests that these populations are yet to be adequately addressed in terms of summarizing the evidence that may have help shape local vascular preventive policies and programs. Consequently, our findings reinforce the positive effects of lifestyle interventions and their significant potentials in formulating cost-beneficial preventive strategies in resource-poor countries.

In conclusion, available data indicate that lifestyle-related interventions are effective in lowering blood pressure in LMIC. The potential of cost-effective, easily scalable lifestyle interventions is attractive in resource-poor settings as a complementary approach to help shape preventive guidelines. Nonetheless, further investigations with sufficient power and scientific rigor would be required to more reliably quantify these effects.



**Contributions** CPB and OHF contributed to conception, study design, literature search, data collection, data analysis, interpretation and writing the manuscript. MO and JOY contributed to literature search, data collection, analysis and writing the manuscript. RC contributed to study design, data analysis, interpretation and critically reviewed the manuscript. ABL,SKLD,NC and SS contributed to literature search, interpretation, analysis and critically reviewed the manuscript. AS, TTWH, RFP, EV,JFF, and JRFN contributed to literature search, data interpretation and critically reviewed the manuscript.

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**Competing interests** - The authors have read the journal's policy and have no Competing Interest.

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**Table 1 - Characteristics of trials identified**

Author	Year	Location	N	Women %	Mean Age(sd)	Population Baseline	Intervention	Control	Duration months
<b>Physical Activity</b>									
Arora. E. <sup>44</sup>	2009	India	20	40	49.6 (5.2)	Diabetes	Resistance training	regular care	2
Barroso.W.K.S. <sup>59</sup>	2008	Brazil	35	80	66.5 (4)		Aerobic . strength and education on lifestyle	education on lifestyle	6
Bündchen. D. C. <sup>78</sup>	2009	Brazil	111	66.6	58 (8.9)	BMI>30	Aerobic and resistance training	no physical activity	3
Kanegusuku. H. <sup>62</sup>	2011	Brazil	24	28	63		Resistance training	no change in lifestyle	4
Jiang. X. <sup>65</sup>	2007	China	167	70	62.11 (7.4)	Angina . MI	Aerobics and resistance	regular care	3
Lamina. S. <sup>72</sup>	2011	Nigeria	245	0	58.63 (7.2)	Hypertension	Interval cycling training	no training	2
McCaffrey. R. <sup>69</sup>	2005	Thailand	54	65	56.7	Hypertension	Yoga	regular care	2
de Meirelles. L. R. <sup>39</sup>	2009	Brazil	19	63.1	49	Hypertensive. CVD. diabetes	Aerobics	no physical activity	3
Monteiro. L. Z. <sup>43</sup>	2010	Brazil	22	100	61 (9.1)		Aerobics	education on lifestyle 1/week	3.25
Terra. D. F. <sup>37</sup>	2008	Brazil	46	100	66.8 (5.6)	Diabetes. osteoporosis. hyper	Resistance exercise training sessions	no change lifestyle routine	3
Thomas. G. N. <sup>67</sup>	2005	China	142	45	69.1 (3.2)		Tai chi	regular level of physical	12
Tsai. J. C. <sup>56</sup>	2003	China	23	50	49.6 (9.3)		Tai chi	remained sedentary	3
Vianna. M. V. A. <sup>76</sup>	2011	Brazil	70	65.7	69.8 (8.5)		Walking. .hydro gym. strengthening. stretching	no change in lifestyle routine	4



Wu. T. Y. <sup>84</sup>	2007	China	36	100	49.7 (6.1)		Aerobics	no lifestyle change	2
<b>Multiple Interventions</b>									
Goldhaber-Fiebert. J. D. <sup>85</sup>	2003	Costa-Rica	61	78.6	50.8 (9.3)	Diabetes. hypertension. hypercholesterolemia	Nutrition education and walking group walking group	Regular care	3
Kisioglu. A N <sup>33</sup>	2003	Turkey	400	100	34.1 (8.6)		Health training and leaflets	Leaflets	6
Lee. L-L. <sup>43</sup>	2007	China	184	41.6	71.3 (6.4)		Walking and public health nurse support	Regular care	8
Marin. G.H. <sup>77</sup>	2009	Argentina	700	65.8	70 (8.3)		Supervised physical activity. dance classes. education on nutrition	Regular care	12
Muda. S. H. <sup>60</sup>	2006	Malaysia	91	0	45 (6.8)		Advice on lifestyle and physical activity	Regular advice on lifestyle	6
Mujica.V. <sup>54</sup>	2010	Chile	51	58.8	51.1 (5.3)		Nutrition education and physical activity	No lifestyle changing program	4.5
Senuzun.F. <sup>53</sup>	2004	Turkey	60	10	54.7 (7.8)	CHD	Home based exercise program with lifestyle advice	Regular care	3
Snehalatha.C. <sup>64</sup>	2008	India	232	20.7	45.9 (5.9)	Hypertension	General lifestyle advice	Standard care	30
Yen.L.L <sup>73</sup>	1996	China	359	35.8	54.5 (14.5)		General lifestyle advice	No self learning packages	2
Pazoki. R. <sup>52</sup>	2007	Iran	335	100	39.4		Physical activity and education material. home visits	Regular care	2

### Dietary Modification

Azadbakht. L. <sup>48</sup>	2005	Iran	55	70.6	41.2 (12.4)	MetS	DASH diet	no diet	6	
Charlton. K. E. <sup>66</sup>	2008	South Africa	80	83.7	61.8 (6.6)		Na + modified food plus salt replacement and fermented milk	foods without modification	2	
China S.C.G. <sup>51</sup>	Salt	2007	China	608	56	59 (10)	Hypertension . CVD. diabetes	Salt substitution for 65% sodium. 25% potassium. 10% magnesium	regular salt	12
He. J. <sup>70</sup>		2006	China	302	53.3	51.4 (9.2)	Hypertension	40g soy bean protein/day	placebo carbohydrate	3
Hu. J <sup>36</sup>		2009	China	192	59	59 (10)	Diabetes	salt substitution for 65% sodium. 25% potassium. 10% magnesium	Regular salt	12
Lin. W.Y <sup>28</sup>		2009	China	132	65.9	34.2 (9.8)	Obesity	very low calorie diet (450 cal/day)	low calorie diet (800 cal/day)	3
Mark. S. D. <sup>31</sup>		1996	China	3318	56	54		Linxian vitamin/mineral supplement/day	lookalike placebo	72
Mukuddem-Petersen. J. <sup>63</sup>		2007	South Africa	43	54.7	45		Walnut diet	regular diet	2
Pereira. M.A.G. <sup>57</sup>		2005	Brazil	22	85.7	45.4 (13.2)		Salt substitution for 50%potassium	regular salt	3
Radhakrishnan. G. <sup>55</sup>		2009	India	85	100	48 (5.4)		25 g rich soy protein (75 mg powder)	placebo of milk	6
Simão. A.N.C. <sup>68</sup>		2010	Brazil	30	100	45.9 (9.8)		25g/day soy. protein	regular diet	3
Sun. J. <sup>45</sup>		2008	China	150	28	51(1)	Diabetes . MetS	Education. dietitian consultation and low glycemic meal replacement	education	6

Torres. M. G. <sup>74</sup>	2010	Brazil	39	90	37.9	BMI>30	Low calorie diet(800 kcal/day). high calcium intake	low calorie diet (800 kcal/day) and low calcium intake	4
Toscani. M. K. <sup>71</sup>	2011	Brazil	26	100	29.4 (5.7)	polycystic ovary syndrome	High protein diet	normal protein diet	2
Wu. H. <sup>47</sup>	2010	China	189	44.2	48.5 (8)	MetS	Flaxseed bread	lifestyle counseling	3
Zhang. G. <sup>46</sup>	2011	China	202		49.8 (7.1)	MetS	Cooked brown rice	white rice	4
<b>Behavioral Counselling</b>									
Figar. S. <sup>30</sup>	2004	Argentina	500	64.8	73	Hypertension	Education and organizational changes in the health	Education and regular care	12
Grace. J. M. <sup>61</sup>	2009	South Africa	122	0	41.7 (8.0)		General lifestyle advice and physical activity	Physical activity	8
Hacihasanoglu.R. <sup>49</sup>	2011	Turkey	80	51.6	58 (8.9)		General lifestyle and drug adherence advice	Education on drug treatment	9
Hammad. E. A. <sup>38</sup>	2011	Jordan	199	64	56 (9.6)	Hypertension. diabetes	General lifestyle advice and medical adherence advice	Regular care	6
Hsieh. C. <sup>75</sup>	2008	China	268	47.7	55.5 (11.7)		General lifestyle advice	Regular care	6
Lin. H. H. <sup>40</sup>	2010	China	73	12.3	61.8 (10.9)	High cholesterol. hypertension. diabetes	General lifestyle advice	Regular care	3
Mendivil. C. O.	2006	Colombia	75	50	51.35 (2.3)				
Metintas.S. <sup>42</sup>	2009	Turkey	498	46.2	57.112.0)	Hypertension. diabetes . CVD	brochure and prescription of lifestyle measures	Brochure	12

Singh. R. B <sup>35</sup> .	1996	India	463	9.5	46 (9.6)	CHD. diabetes. hypertension	fat modified. fruit and vegetable enriched diet plus daily moderate exercise	Education	36
Singh. R. B. <sup>58</sup>	1993	India	621	9.5	46.3 (8.6)	CHD. diabetes. hypertension	Nutrition education and walking sessions	Education	6
Uzcategui .C D.	1999	Venezuela	20	0	N.I.	Hypertension	Diet. exercise tapes. relaxation. education	Regular care	5

Sd- Standard deviation, CVD- Cardiovascular disease, CHD- Coronary Heart Disease, MetS- Metabolic Syndrome, MI – Myocardial infarction, N.I. – Non informed

**Table 2- Subgroup analysis for physical activity interventions effect on blood pressure in LMIC.**

	N Studies	Systolic Blood Pressure			Diastolic Blood Pressure		
		WMD	CI 95%	pvalue	WMD	CI 95%	pvalue
Drugs				0.378			0.000
Yes	7	-12.13	(-18.64, -5.62)		-7.38	(-12.27, -2.5)	
No	7	-10.5	(-18.18,-2.83)		-5.2	(-8.42, -1.98)	
Ethnicity				0.000			0.002
Asian	6	-11.02	(-20.12, -1.93)		-6.27	(-12.07, -0.47)	
African	1	-13.09	(-13.09,-9.97)		-3.3	(-5.21, -1.39)	
South American	7	-11.82	(-11.82, -6.49 )		-7.38	(-11.14, -3.63)	
Duration				0.000			0.177
< 4 months	10	-13.8	(-19.3, -8.31)		-7.35	(-10.91, -3.78)	
4-12 months	4	-5.39	(-12.46, 1.68)		-4.61	(-10.96, 1.75)	
Sample Size				0.000			0.000
<30	5	-12	(-23.38, -0.62)		-7.48	(-13.01, -1.96)	
30-100	5	-13.28	(-20.17, -6.39)		-9.08	(-15.01, -3.14)	
>100	4	-8.15	(-15.77, -0.53)		-2.77	(-6.13, 0.59)	
Intensity METs				0.000			0.623
≤ 6 METs	9	-10.77	(-16.87, -4.66)		-6.42	(-10.64, -2.2)	
>6METs	5	-13.11	(-19.67, -6.56)		-6.75	(-11.05, -2.45)	
Type activity				0.088			0.001
Aerobic and resistance	7	-12.9	(-20.9, -5.44)		-8.01	(-13.04, -2.98)	
Aerobic	4	-8.52	(-14.2,-2.85)		-2.33	(-4.28, -0.37)	
Resistance	3	-13.02	(-27.25, 1.21)		-9.78	(-15.9, -3.66)	
Weekly hours				0.000			0.000
≤ 2 hours	4	-1.89	(-4.27, 0.48)		-2.02	(-5.05, 1.05)	
2 to 3 hours	10	-14.78	(-19.23, -10.33)		-8.43	(-12.17, -4.68)	
Co-morbidities				0.000			0.547
No	6	-8.75	(-16.47,-1.02)		-5.4	(-9.63,-1.17)	
Yes	8	-13.26	(-19.35,-7.16)		-7.47	(-11.89,-3.05)	
Randomization				0.068			0.001
Yes	5	-12.83	(-21.70,-3.97)		-6.18	(-12.00,-0.36)	
No	4	-13.04	(-19.93,-6.15)		-6.20	(-10.29,-2.11)	
Unclear	5	-8.38	(-15.77,-0.99)		-12.06	(-12.06,-2.61)	
Blinding				0.097			0.004
Yes	2	-12.62	(-33.71,8.48)		-4.61	(-12.59,3.37)	
No	2	-8.74	(-13.84,-3.63)		-2.96	(-8.17,2.26)	
Unclear	10	-11.76	(-17.53,-6.00)		-7.79	(-11.75,-3.83)	
Intention to				0.000			0.000
Yes	4	-7.74	(-16.56,1.08)		-2.33	(-5.86,1.21)	
No	10	-13.27	(-17.60,-8.95)		-8.01	(-11.51,-4.50)	

WMD- Weighted Mean Difference in mm Hg, METs – Metabolic equivalent

\* pvalues for heterogeneity between groups

**Table 3 - Subgroup analysis for behavioral counseling interventions effect on blood pressure in LMIC.**

	N Studies	Systolic Blood Pressure			Diastolic Blood Pressure		
		WMD	CI 95%	pvalue	WMD	CI 95%	pvalue
Drugs				0.000			0.138
Yes	2	-9.25	(-17.25,-		-2.93	(-5.35,-	
No	5	-0.73	(-3.10,1.63)		-0.89	(-2.97,1.19)	
Duration				0.000			0.000
< 6 months	4	-5.84	(-9.60,-		-2.81	(-5.32,-	
6 to 30 months	3	0.96	(-1.56,3.48)		0.22	(-0.81,1.25)	
Sample Size				0.116			0.604
<200	3	-5.12	(-		-1.84	(-4.25,0.56)	
200-498	4	-1.39	(-3.77,1.00)		-1.07	(-3.52,1.38)	
Co-morbidities				0.572			0.009
No	3	-1.62	(-5.32,2.08)		-0.77	(-2.53,0.99)	
Yes	4	-3.83	(-9.02,1.36)		-1.76	(-5.00,1.48)	

WMD- Weighted Mean Difference in mm Hg

\* pvalues for heterogeneity between groups

**Table 4 - Subgroup analysis for dietary modifications on blood pressure in LMIC.**

Type of diet	N Studies	Systolic Blood Pressure			Diastolic Blood Pressure		
		WMD	CI 95%	pvalue	WMD	CI 95%	pvalue
				0.000			0.009
Mineral replacement based	3	-6.90	(-8.6.-5.2)		-1.90	(-4.50. 0.70)	
Grains and fiber	1	-1.47	(-3.9.-1.01)		-0.10	(-1.90. 1.77)	
Protein Rich	4	0.93	(-4.6. 6.46)		-3.2	(-7.50. 1.10)	
Nuts and seeds	2	-1.57	(-4.73.1.59)		-0.49	(-2.42. 1.43)	
Complex Dietary Pattern	3	-3.47	(-5.45.-1.5)		-2.25	(-3.60.-0.89)	

WMD- Weighted Mean Difference in mm Hg

\* pvalues for heterogeneity between groups

**Table 5- Subgroup analysis for multiple interventions effect on blood pressure in LMIC.**

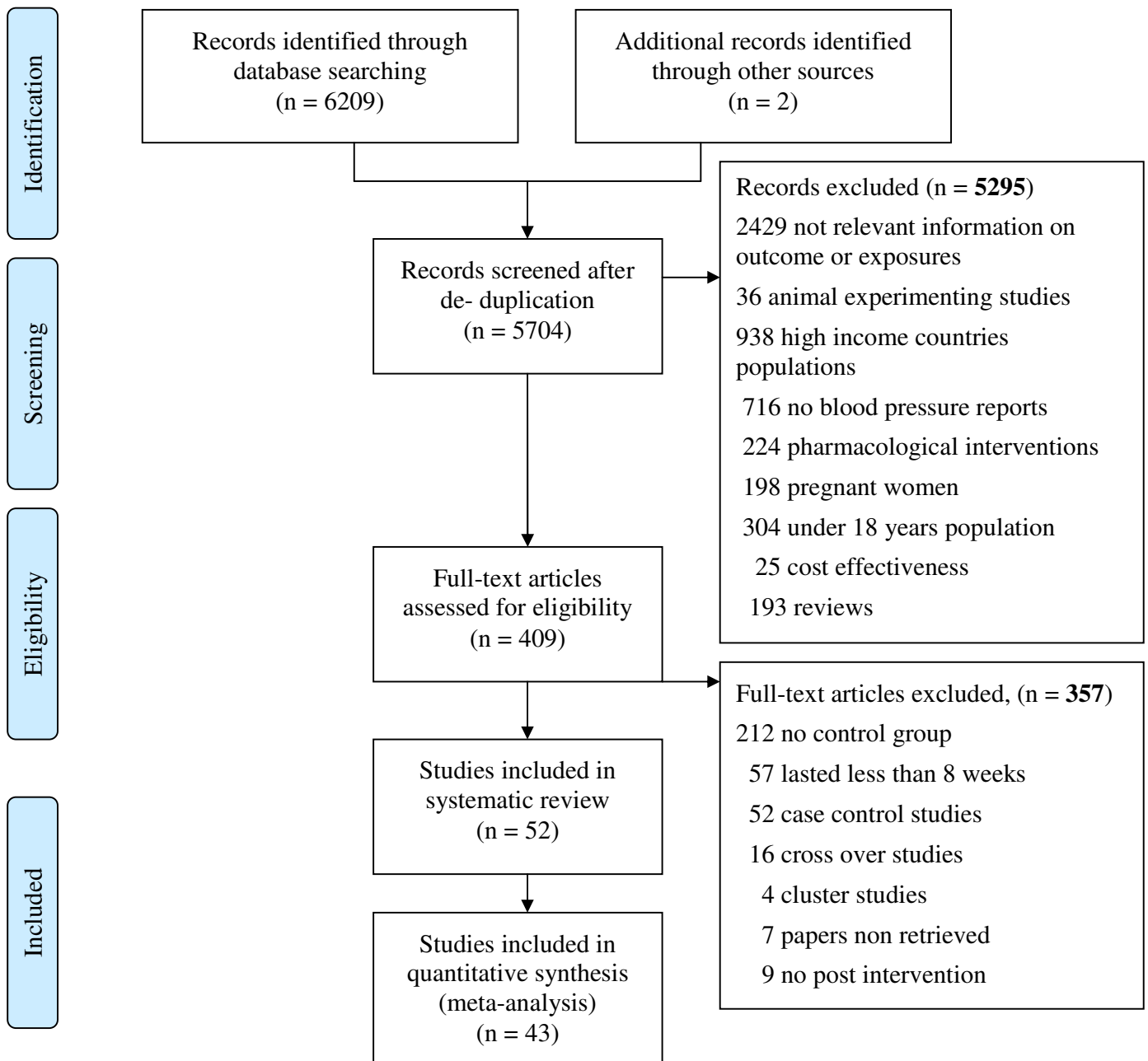
	N Studies	Systolic Blood Pressure			Diastolic Blood Pressure		
		WMD	95% CI	pvalue*	WMD	95% CI	pvalue*
Delivery				0.115			0.889
Diet/Physical Activity	4	-4.79	(-8.74,-0.83)		-2.07	(-3.54,-0.61)	
Behavioral							
counseling/Physical	4	-7.16	(-12.17,-2.15)		-2.86	(-5.4,-0.32)	
Ethnicity				0.000			0.225
Asian	5	-7.75	(-10.72,-4.78)		-2.22	(-3.94,-0.5)	
American	3	-3.01	(-3.01,-0.37)		-2.80	(-4.57,-1.03)	
Duration				0.007			0.639
< 6 months	4	-6.22	(-12.19,-0.26)		-4.05	(-7.67,-0.42)	
6 to 12 months	4	-5.49	(-8.84,-2.14)		-1.70	(-2.67,-0.72)	
Co-morbidities				0.194			0.575
No	4	-6.44	(-11.95,-0.93)		-2.10	(-4.02,-0.18)	
Yes	4	-5.46	(-9.09,-1.83)		-2.93	(-5.08,-0.78)	
Blinding outcome				0.738			0.504
assessor							
Yes	5	-6.86	(-9.18,-4.54)		-3.34	(-5.67,-1.01)	
Unclear	3	-6.17	(-12.82,0.48)		-1.72	(-3.27,-0.17)	
Intention to treat				0.887			0.789
Yes	4	-5.79	(-9.15,-2.43)		-2.10	(-3.72,-0.49)	
No	4	-6.16	(-12.02,-0.30)		-2.89	(-5.28,-0.50)	

WMD- Weighted Mean Difference in mm Hg

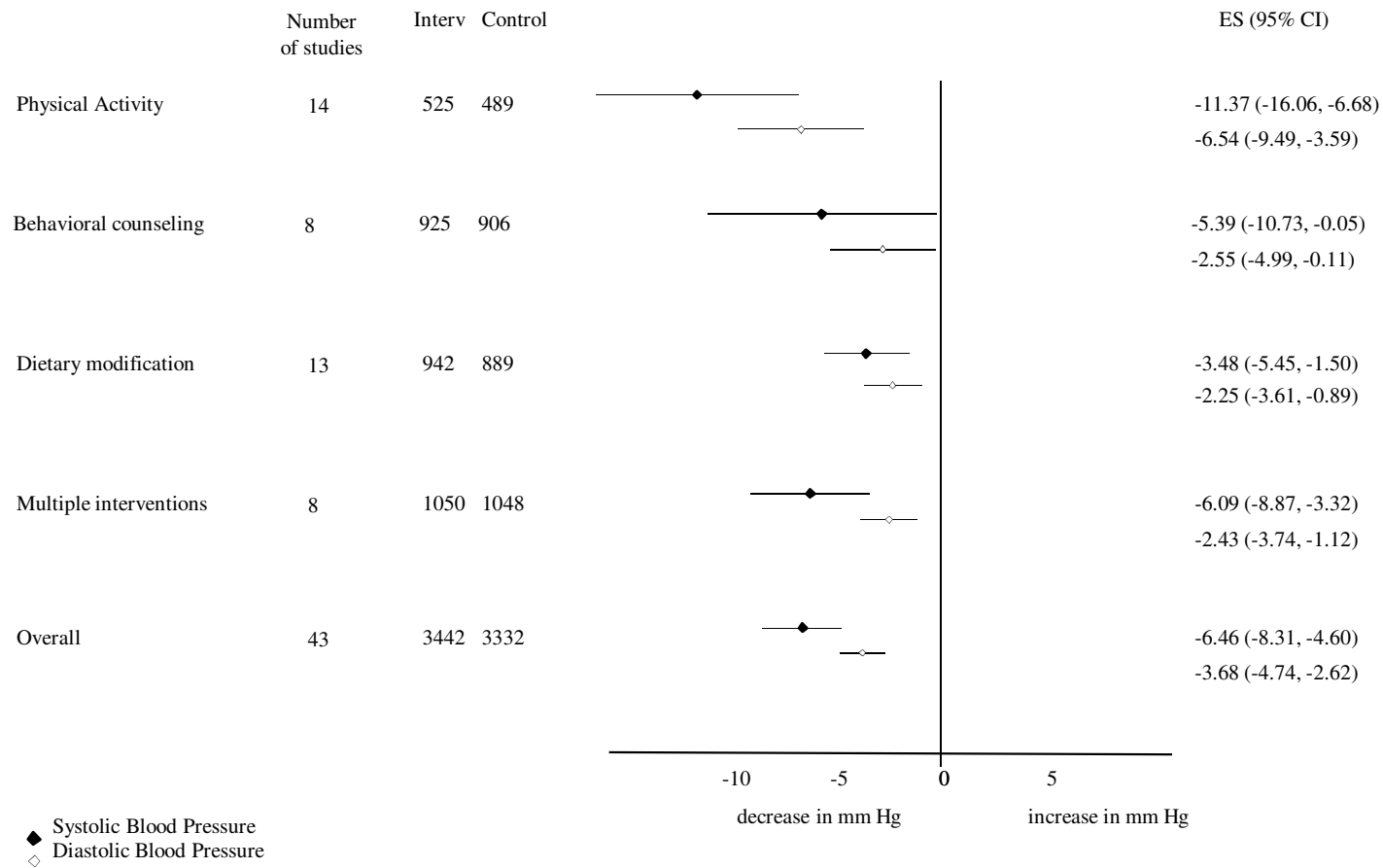
\* pvalues for heterogeneity between groups



Figure 1 – Study flow diagram



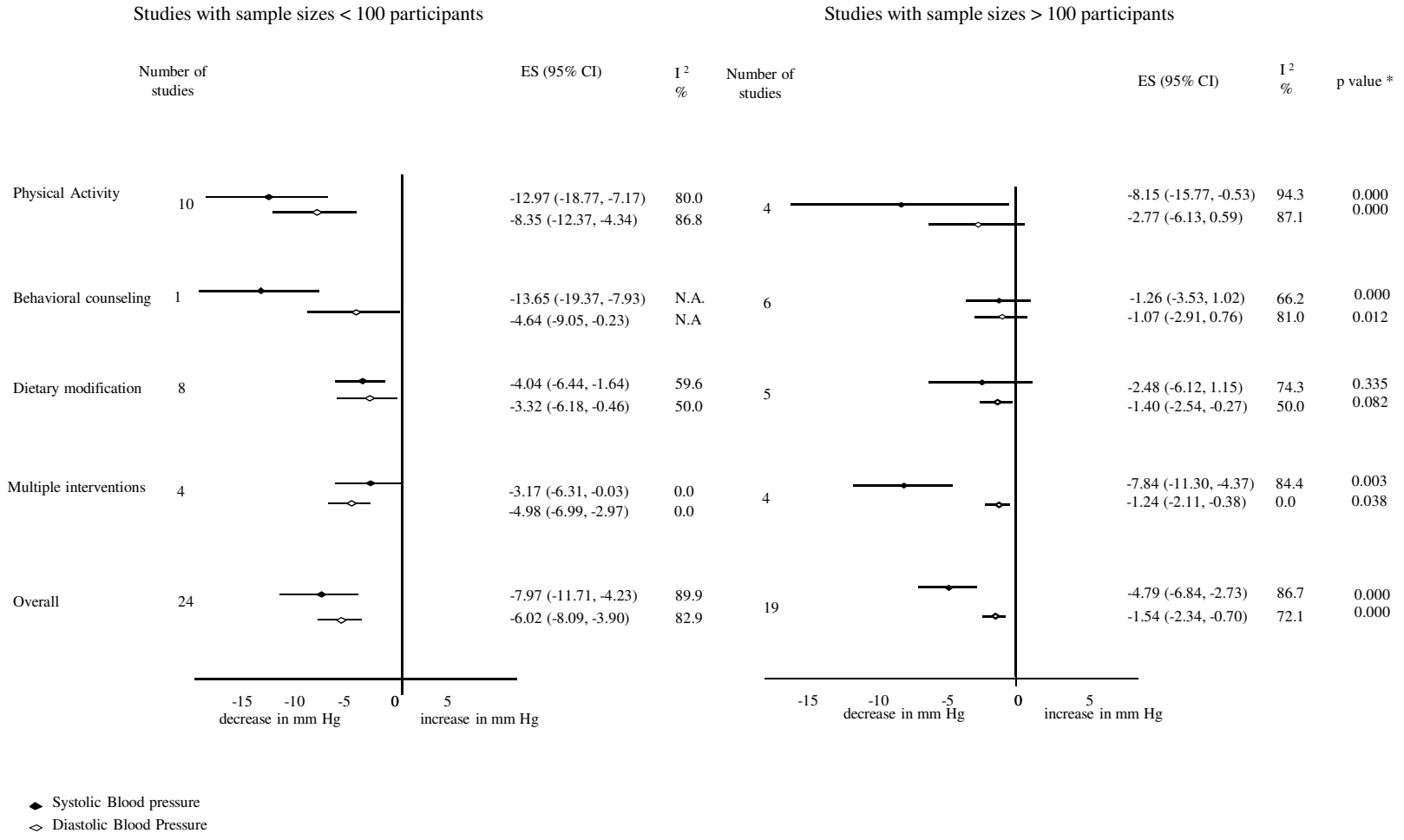




ES- Effect Size, CI- Confidence Interval 95%

Heterogeneity ( $I^2$ , pvalue) within groups : Behavioral counseling (94.9%,  $p < 0.001$ ; 90.3%,  $p < 0.001$ ), Physical Activity (88.7%,  $p = 0.000$ ; 89.5%,  $p < 0.001$ ), Dietary Modification (64.5%,  $p = 0.000$ ; 68.5%,  $p < 0.001$ ), Multiple interventions (75.9%,  $p = 0.000$ ; 52.9%,  $p = 0.038$ ), Overall (89.0%,  $p = 0.000$ ; 85.4%,  $p < 0.001$ ) to systolic and diastolic pressure respectively.

**Figure 2 - Lifestyles interventions effect on blood pressure in LMIC populations**



**Figure 3- Lifestyle interventions effect on blood pressure in LMIC populations by sample sizes**

**Low income:** Afghanistan. Bangladesh. Benin. Burkina Faso. Burundi. Cambodia. Central African Republic. Chad. Comoros. Côte d'Ivoire. Democratic People's Republic of Korea. Democratic Republic of the Congo. Eritrea. Ethiopia. Gambia. Ghana. Guinea. Guinea-Bissau. Haiti. India. Kenya. Kyrgyzstan. Lao People's Democratic Republic. Liberia. Madagascar. Malawi. Mali. Mauritania. Mongolia. Mozambique. Myanmar. Nepal. Niger. Nigeria. Pakistan. Papua New Guinea. Rwanda. Sao Tome and Principe. Senegal. Sierra Leone. Solomon Islands. Somalia. Sudan. Tajikistan. Timor-Leste. Togo. Uganda. United Republic of Tanzania. Uzbekistan. Viet Nam. Yemen. Zambia. Zimbabwe

**Lower middle income:** Albania. Algeria. Angola. Armenia. Azerbaijan. Belarus. Bhutan. Bolivia. Bosnia and Herzegovina. Cameroon. Cape Verde. China. Colombia. Congo. Cuba. Djibouti. Dominican Republic. Ecuador. Egypt. El Salvador. Fiji. Georgia. Guatemala. Guyana. Honduras. Indonesia. Iran (Islamic Republic of). Iraq. Jamaica. Jordan. Kiribati. Lesotho. Maldives. Marshall Islands. Micronesia (Federated States of). Morocco. Namibia. Nicaragua. Paraguay. Peru. Philippines. Republic of Moldova. Samoa. Sri Lanka. Suriname. Swaziland. Syrian Arab Republic. Thailand. The former Yugoslav Republic of Macedonia. Tonga. Tunisia. Turkmenistan. Ukraine. Vanuatu

**Upper middle income:** Argentina. Belize. Botswana. Brazil. Bulgaria. Chile. Costa Rica. Croatia. Dominica. Equatorial Guinea. Gabon. Grenada. Hungary. Kazakhstan. Latvia. Lebanon. Libyan Arab Jamahiriya. Lithuania. Malaysia. Mauritius. Mexico. Montenegro. Oman. Palau. Panama. Poland. Romania. Russian Federation. Saint Kitts and Nevis. Saint Lucia. Saint Vincent and the Grenadines. Serbia. Seychelles. Slovakia. South Africa. Turkey. Uruguay. Venezuela (Bolivarian Republic of)

**High income:** Andorra. Antigua and Barbuda. Australia. Austria. Bahamas. Bahrain. Barbados. Belgium. Brunei Darussalam. Canada. Cyprus. Czech Republic. Denmark. Estonia. Finland. France. Germany. Greece. Iceland. Ireland. Israel. Italy. Japan. Kuwait. Luxembourg. Malta. Monaco. Netherlands. New Zealand. Norway. Portugal. Qatar. Republic of Korea. San Marino. Saudi Arabia. Singapore. Slovenia. Spain. Sweden. Switzerland. Trinidad and Tobago. United Arab Emirates. United Kingdom. United States of America Cook Islands. Nauru. Niue and Tuvalu are not categorized into income groups and are therefore excluded from the computation of aggregate indices by income group.

Appendix 2 – Methodological Supplement – Search Strategy on Pubmed

Population	(Developing Countr*[tw] OR Under Developed countr*[tiab] OR UnderDeveloped countr*[tiab]OR less Developed countr*[tiab] OR Developing nation*[tiab] OR Under Developed nation*[tiab] OR UnderDeveloped nation*[tiab] OR less Developed nation*[tiab] OR Third World*[tiab] OR low resource countr*[tiab] OR low resource nation*[tiab] OR africa[mesh] OR (africa*[tiab] NOT african americans[tiab]) OR South America[mesh] OR South America*[tiab] OR latin America*[tw] OR central America[mesh] OR ((asia[mesh] OR asia*[tiab]) NOT japan*[mesh]))
Intervention	(life style*[tw] OR lifestyl*[tw] OR diet therapy[mesh] OR Sodium Restrict*[tiab] OR salt Restrict*[tiab] OR low Sodium*[tiab] OR low salt*[tiab] OR Potassium, Diet* [tw]OR Magnesium [tw]OR Calcium [tw] OR fat Restrict*[tiab] OR low fat*[tiab] OR Carbohydrate Restrict*[tiab] OR low carb*[tiab] OR Caloric Restrict*[tw] OR Food, Formulated[tw] OR Formulated Food*[tw] OR diet[tw] OR dietary[tw] OR weight loss*[tw] OR losing weight[tw] OR Weight Reduction*[tiab] OR Disease Management*[tw] OR Exercise[mesh] OR Exercise therapy[mesh] OR Exercise test[mesh] OR Exercise Movement Techniques[mesh] OR kinesiotherap*[tw] OR Physical Endurance[mesh] OR Anaerobic*[tiab] OR aerobic*[tiab] OR Exercise*[tiab] OR Resistance Training*[tiab] OR Motor activit*[tw] OR Physical Activit*[tiab] OR Locomotor Activit*[tiab] OR social support*[tw] OR Social Network*[tiab] OR Tobacco Use Cessation*[tw]OR Smoking cessation*[tw] OR Alcohol Drink*[tw] OR Alcohol consum*[tw] OR Drinking Alcohol*[tw] OR Alcoholi*[tw] OR non pharmacol*[tw] OR relaxation therap* [tw] OR tai-ji [tw] OR yoga [tw])
Outcome	(Blood pressure*[tiab] OR Hypertension[mesh:noexp] OR Hypertension[tiab] OR Systolic Pressure[tiab] OR Diastolic Pressure [tiab] OR Pulse Pressure [tiab] )
Study Type	(Clinical Trial[pt] OR Randomized Controlled Trial[pt] OR controlled clinical trial [pt] OR randomized [tiab] OR placebo [tiab] OR randomly [tiab] OR trial [tiab] OR groups [tiab] OR Comparative Study[pt] OR Compar*[tiab] OR Clinical Trial*[tiab] OR Evaluation Studies[pt] OR Evaluat*[tiab] OR Effectiv*[tiab] OR utilit*[tiab] OR Validation Studies[pt] OR Validat*[tiab] OR reliab*[tiab] OR relevan*[tiab])

*Appendix 3 – Methodological Supplement – Missing Variance imputation*

The mean difference variance estimation was calculated according different information reporting across trials and it was imputed whenever standard deviation, standard errors or confidence interval 95% was reported for baseline and/or after intervention measurements. Variance imputation to continuous variables response seems to be a field of controversy <sup>86</sup> but some studies have shown that different methods won't alter results of meta-analysis given the appropriate attention to each trial <sup>87 88</sup>. In fact, a number of meta-analysts have to deal with a trials lacking variance of effect estimates <sup>89</sup>.

Since only 6 interventions reported variance of mean difference between pre and post intervention in each group, we imputed variance separately in each trial according to available information in each study.

When the study would inform the p value of the mean change variance we calculated the standard deviation. When the study informed t statistics of the mean change and the p value, the t statistics were used. From the sd baseline, final and change we calculated the Correlation Coefficient and used its median from the available ones. We then calculated the median from the studies that provided SD change .

When we had two datasets, one using imputation of variance with assumed correlation coefficient from studies that reported means and standard deviations for systolic and diastolic pressure at baseline (pre intervention) and final (post intervention) <sup>86 88</sup> and another dataset substituting missing variance information using the median of available standard deviations reported in other studies <sup>86</sup>. The formula used to derive Correlation Coefficient from informed variance in other studies is shown (*Supplementary Appendix 2*). A conservative approach was used to input variance from other studies Correlation Coefficient. We analysed available sd descriptively and ranged from the median (0,53) to the maximum (0,87) according to the critical p values informed in the studies (e.g.  $p < 0,001$  or  $p < 0,05$ ).

Table 1- Equations to calculate mean difference and effect size

Standard deviation from Standard Error	$SD = \sqrt{n} SE$
Standard deviation from a Confidence Interval 95%	$SD = \frac{\sqrt{n} (\text{upper limit} - \text{lower limit})}{3.92}$
Standard deviation of the mean change from p value or t statistics	$SD_{\text{change}} = \frac{\sqrt{n} \text{Mean}_{\text{change}}}{t_{n-1; p \text{ value}}}$
Correlation Coefficient	$\text{Corr} = \frac{SD_{\text{baseline}}^2 + SD_{\text{final}}^2 - SD_{\text{change}}^2}{2 SD_{\text{baseline}} SD_{\text{final}}}$
Standard deviation from Correlation Coefficient	$SD_{\text{change}} = \sqrt{SD_{\text{baseline}}^2 + SD_{\text{final}}^2 - (2 \text{ Corr} SD_{\text{baseline}} SD_{\text{final}})}$



Appendix 4 – PRISMA checklist for systematic review and meta analysis.

Section/topic	#	Checklist item	Reported on page #
<b>TITLE</b>			
Title	1	Identify the report as a systematic review, meta-analysis, or both.	1
<b>ABSTRACT</b>			
Structured summary	2	Provide a structured summary including, as applicable: background; objectives; data sources; study eligibility criteria, participants, and interventions; study appraisal and synthesis methods; results; limitations; conclusions and implications of key findings; systematic review registration number.	2
<b>INTRODUCTION</b>			
Rationale	3	Describe the rationale for the review in the context of what is already known.	3
Objectives	4	Provide an explicit statement of questions being addressed with reference to participants, interventions, comparisons, outcomes, and study design (PICOS).	3
<b>METHODS</b>			
Protocol and registration	5	Indicate if a review protocol exists, if and where it can be accessed (e.g., Web address), and, if available, provide registration information including registration number.	4
Eligibility criteria	6	Specify study characteristics (e.g., PICOS, length of follow-up) and report characteristics (e.g., years considered, language, publication status) used as criteria for eligibility, giving rationale.	4
Information sources	7	Describe all information sources (e.g., databases with dates of coverage, contact with study authors to identify additional studies) in the search and date last searched.	4/5
Search	8	Present full electronic search strategy for at least one database, including any limits used, such that it could be repeated.	Appendix 2
Study selection	9	State the process for selecting studies (i.e., screening, eligibility, included in systematic review, and, if applicable, included in the meta-analysis).	Figure 1
Data collection process	10	Describe method of data extraction from reports (e.g., piloted forms, independently, in duplicate) and any processes for obtaining and confirming data from investigators.	4/5
Data items	11	List and define all variables for which data were sought (e.g., PICOS, funding sources) and any assumptions and simplifications made.	5
Risk of bias in individual studies	12	Describe methods used for assessing risk of bias of individual studies (including specification of whether this was done at the study or outcome level), and how this information is to be used in any data synthesis.	5
Summary measures	13	State the principal summary measures (e.g., risk ratio, difference in means).	5
Synthesis of results	14	Describe the methods of handling data and combining results of studies, if done, including measures of consistency (e.g., $I^2$ ) for each meta-analysis.	5/6
<b>Section/topic</b>			
Risk of bias across studies	15	Specify any assessment of risk of bias that may affect the cumulative evidence (e.g., publication bias, selective reporting within studies).	6

Additional analyses	16	Describe methods of additional analyses (e.g., sensitivity or subgroup analyses, meta-regression), if done, indicating which were pre-specified.	6
<b>RESULTS</b>			
Study selection	17	Give numbers of studies screened, assessed for eligibility, and included in the review, with reasons for exclusions at each stage, ideally with a flow diagram.	7/Figure 1
Study characteristics	18	For each study, present characteristics for which data were extracted (e.g., study size, PICOS, follow-up period) and provide the citations.	21-24
Risk of bias within studies	19	Present data on risk of bias of each study and, if available, any outcome level assessment (see item 12).	6
Results of individual studies	20	For all outcomes considered (benefits or harms), present, for each study: (a) simple summary data for each intervention group (b) effect estimates and confidence intervals, ideally with a forest plot.	Figure 2
Synthesis of results	21	Present results of each meta-analysis done, including confidence intervals and measures of consistency.	,25,26,27, 28, Figure 2, Figure 3
Risk of bias across studies	22	Present results of any assessment of risk of bias across studies (see Item 15).	25,26,28, Figure 3
Additional analysis	23	Give results of additional analyses, if done (e.g., sensitivity or subgroup analyses, meta-regression [see Item 16]).	25,26,28, Figure 3
<b>DISCUSSION</b>			
Summary of evidence	24	Summarize the main findings including the strength of evidence for each main outcome; consider their relevance to key groups (e.g., healthcare providers, users, and policy makers).	10,12,13
Limitations	25	Discuss limitations at study and outcome level (e.g., risk of bias), and at review-level (e.g., incomplete retrieval of identified research, reporting bias).	12,13
Conclusions	26	Provide a general interpretation of the results in the context of other evidence, and implications for future research.	12,13
<b>FUNDING</b>			
Funding	27	Describe sources of funding for the systematic review and other support (e.g., supply of data); role of funders for the systematic review.	13

## CONSIDERAÇÕES FINAIS

A doença cardiovascular é a mais importante causa de óbito dentre as doenças crônicas não transmissíveis no mundo. Tanto Curitiba quanto o Brasil não são exceção. Durante a elaboração destes artigos, analisou-se a mortalidade por DCV em nível regional e nacional e, ainda, reuniu-se evidência para o controle de um importante fator de risco nos países em desenvolvimento. Os resultados apresentados nesta tese permitem algumas considerações.

Em Curitiba, a principal causa de óbito cardiovascular na última década foi o Infarto Agudo do Miocárdio. Embora a qualidade de informação nesta capital esteja dentro dos padrões de qualidade da Organização Mundial de Saúde, o declínio na tendência de mortalidade por esta causa, encontrado nos dois gêneros e em todas as faixas etárias deve ser interpretado com cautela. Em primeiro lugar, Curitiba apresentava uma das taxas de óbito mais altas do país no início do estudo. Em segundo lugar, a participação dos setores primário, secundário e terciário no total de óbitos esperados e não observados permanece desconhecida, uma vez que os sistemas de informação entre estes níveis não são compatíveis com o sistema de mortalidade. Adicionalmente, estudos transversais de fatores de risco cardiovascular vem mostrando o aumento de prevalência de obesidade, diabetes, hipertensão, dieta inadequada e inatividade física, colocando em risco a atual tendência de declínio observada em todas as idades e nos dois gêneros em Curitiba. Neste sentido, a uniformização e a compatibilidade de informações entre os diversos setores de saúde possibilitarão estudos longitudinais de estudo de causa para o planejamento de longo prazo do combate a esta causa de óbito que embora declinante, ainda se apresenta como a mais importante nesta cidade.

No Brasil, encontrou-se que a tendência de mortalidade por DIC permaneceu estagnada nos últimos 10 anos. O declínio registrado anteriormente para a tendência de mortalidade por DIC passou a um platô na última década e esta mudança de tendência foi influenciada pelas disparidades nas tendências regionais, explicadas em grande parte pelas mudanças sócio econômicas, vistas no mesmo

período. Adicionalmente, as projeções para os próximos anos, indicam que as inequidades em saúde entre regiões norte/nordeste e regiões sul/sudeste tendem a aumentar, direcionando a tendência geral do país a um aumento no futuro. Estes achados, evidenciam a necessidade iminente de pesquisa local e estratégias adequadas às necessidades regionais para o combate desta importante causa de morbidade e mortalidade no Brasil.

Nos países em desenvolvimento, onde o impacto da doença cardiovascular é o dobro do impacto observado nos países desenvolvidos, um importante fator de risco, a hipertensão permanece subdiagnosticado e não controlado. Nossos resultados indicam que as intervenções de estilo de vida são terapias adjuvantes adequadas para a prevenção e o controle da pressão arterial destas populações e, portanto, deveriam ser priorizadas. Entretanto, a qualidade da pesquisa e das publicações deve ser enfatizada, a fim de produzir evidências com poder suficiente para reprodução destas intervenções em outros subgrupos expostos, e que em última estância, estão vulneráveis à doença cardiovascular.

No conjunto deste trabalho, foram utilizadas metodologias de epidemiologia básica e epidemiologia clínica para a análise abrangente e possíveis mecanismos de prevenção desta importante causa de óbito, a doença cardiovascular. Nossos resultados indicam possibilidades e hipóteses de ações integradas entre pesquisa e serviços de saúde. Por outro lado, foram apresentadas evidências de alternativas adequadas para a prevenção e controle de um importante fator de risco, a hipertensão.

Espera-se, finalmente, que este trabalho contribua para a compreensão e para o enfrentamento da doença cardiovascular em nosso meio e em outras populações.

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